

# KJHP

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Aims and Scope

The *Korean Journal of Health Promotion (KJHP)* is an open access, multidisciplinary journal dedicated to publishing high-quality research in various areas of the medical, nursing, nutritional, physical educational, epidemiological, and public health sciences associated with health promotion and disease prevention. *KJHP*, which has been published continuously since 2001, is an official journal of the Korean Society for Health Promotion and Disease Prevention.

The aim of the *KJHP* is to advance and disseminate new knowledge and scientific information in all the areas associated with health promotion and disease prevention. *KJHP* publishes original articles, narrative reviews, systematic reviews and meta-analyses, letters to the editor, and perspectives in English.

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# Identifying the Optimal Waist Circumference-Based Body Adiposity Index for Detecting Metabolically Obese but Normal-Weight Older Koreans: A Cross-Sectional Study

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## ABSTRACT

**Background:** The concept of being metabolically obese but normal weight (MONW) has emerged as a significant public health concern. Individuals with MONW exhibit insulin resistance and are at elevated risk for chronic diseases, despite having a body mass index (BMI)  $<25 \text{ kg/m}^2$ . Given the high prevalence of MONW among older Asian adults and the role of abdominal visceral fat in metabolic abnormalities in individuals with MONW, identifying an accurate waist circumference (WC)-based body adiposity index for older Asian adults is essential. Thus, this study aimed to determine the most accurate WC-based body adiposity index for identifying MONW in older Korean adults that are not underweight or obese.

**Methods:** This cross-sectional study included 7,846 participants (3,536 males and 4,310 females) aged  $\geq 60$  years with a BMI of  $18.5\text{--}24.9 \text{ kg/m}^2$  from the Korea National Health and Nutrition Examination Survey. MONW was defined using metabolic syndrome criteria and sex-specific triglyceride–glucose Index cutoff points. Associations were analyzed using logistic regression, and discriminative ability was assessed using receiver operating characteristic curves. The analysis included BMI, WC and various WC-based adiposity indices, including the visceral adiposity index (VAI), body roundness index, waist-to-height ratio, conicity index, weight-adjusted waist index, and a body shape index.

**Results:** While all WC-based indices were associated with MONW, only the VAI showed a stronger association (odds ratios [ORs], 26.542 in males and 21.495 in females) and superior discriminative ability (area under the curves [AUCs], 0.826 in males and 0.802 in females) in both sexes compared with WC (ORs, 4.538 and 4.060; AUCs, 0.679 and 0.676). Other indices showed weaker associations and lower areas under the curves than did the VAI and WC, regardless of sex.

**Conclusions:** The VAI may serve as the most effective indicator for identifying MONW among older Korean adults that are not underweight or obese.

**Keywords:** Body adiposity index, Metabolically obese but normal weight, Older adults

## INTRODUCTION

Recently, a new concept of obesity, “metabolically obese but

normal weight (MONW)”, has gained attention as a significant public health concern [1,2]. MONW individuals are characterized by insulin resistance and are highly vulnerable to chronic

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diseases, despite having a body mass index (BMI)  $<25 \text{ kg/m}^2$  [1,3]. Notably, there are substantial numbers of MONW in older Asian adults with type 2 diabetes, cardiovascular disease, or other chronic health concerns [4,5]. The incidences of cardiovascular disease and all-cause mortality for these individuals are even higher than those of their metabolically healthy counterparts with a BMI  $\geq 25 \text{ kg/m}^2$  [6,7]. These findings underscore the limitations of BMI, which is a measure of body weight relative to height without considering fat distribution. Many previous reports have demonstrated the need for a more precise indicator of MONW in older Asian adults who do not meet BMI-based obesity criteria to prevent and manage health concerns related to metabolic abnormalities.

Since the accumulation of visceral fat in the abdominal region is considered a key contributor to metabolic abnormalities in MONW individuals, evaluating abdominal obesity is crucial. For years, waist circumference (WC) has been used to evaluate abdominal obesity [8,9]. However, WC is affected by differences in height and overall body size, making it less effective for evaluating abdominal adiposity [10,11]. To address the limitations of WC, several WC-based body adiposity indices, including the visceral adiposity index (VAI), body roundness index (BRI), waist-to-height ratio (WHtR), conicity index (CI), weight-adjusted-waist index (WWI), and a body shape index (ABSI), have been suggested [12–16]. All of these body adiposity indices have been reported as distinguishing indicators of various health concerns [12–16].

A cross-sectional study, involving Brazilian adults aged 20–59 years old, conducted by Ferreira et al. [17] has reported that the VAI showed a better discriminative ability of MONW, compared to BMI, WHtR, waist-to-hip ratio, and waist and neck circumference. However, given the different characteristics in anthropometry, eating and physical activity habits between Europeans and Koreans, it remains uncertain whether the VAI is an equally effective indicator of MONW for older Asian adults. Additionally, the relationship between MONW and these adiposity indices in older Korean adults who are not underweight or obese ( $18.5 \leq \text{BMI} < 25.0 \text{ kg/m}^2$ ) has not yet been studied. Therefore, this study aimed to investigate the associations between WC-based adiposity indices and MONW in older Korean adults. Furthermore, we compared the performance of these indices in discriminating MONW.

## METHODS

### Study design and participants

This cross-sectional study analyzed data from the Korea National Health and Nutrition Examination Survey (KNHANES) collected between 2014 and 2022. The KNHANES is an ongoing, nationally representative surveillance system established in 1984 to assess the health, nutritional status, and lifestyle behaviors of the Korean population. From the initial dataset, 7,846 individuals (3,536 males and 4,310 females) aged  $\geq 60$  years with a BMI between 18.5 and  $24.9 \text{ kg/m}^2$  were included. All participants provided written informed consent. This study was approved by the Institutional Review Board of Changwon National University (approval number: 7001066-202404-HR-056) and conducted in accordance with the principles of the Declaration of Helsinki.

### Measurements and body adiposity indices

Height was measured to the nearest 0.1 cm using a stadiometer with participants standing barefoot. Weight was recorded to the nearest 0.1 kg using a calibrated digital scale while the participants wore light clothing. BMI was obtained using the following formula:  $\text{weight (kg)}/\text{height (m)}^2$ . WC was measured to the nearest 0.1 cm using a nonelastic glass fiber tape. Blood pressure was manually measured three times using a standard sphygmomanometer in a mobile examination unit, and the mean of the three readings was used in the analysis. Venous blood samples were drawn in the morning following an overnight fast of more than 8 hours. Serum concentrations of glucose, triglycerides, and high-density lipoprotein (HDL) cholesterol were determined using enzymatic or homogeneous enzymatic colorimetric methods with a Hitachi 7600–210 automatic analyzer (Hitachi). WC-based body adiposity indices, including the VAI, BRI, WHtR, CI, WWI, and ABSI, were computed as described in Table 1.

### Evaluation of metabolically obese but normal weight and metabolically healthy and normal weight

In the current study, participants were classified as MONW if they met either of the two criteria based on the method proposed by Lee et al. [18]: the presence or absence of metabolic syndrome (MS), or exceeding the sex-specific triglyceride glucose (TyG) index cutoff point. Those who met neither criterion were classified as metabolically healthy and normal weight (MHNW) individuals. Because MS encompasses a constellation



**Table 1.** Formulas of WC-based body adiposity indices

Index	Formula
Visceral adiposity index	Males: $(WC/[39.68+(1.88 \times BMI)]) \times (TG/1.03) \times (1.31/HDL-C)$ Females: $(WC/[36.58+(1.89 \times BMI)]) \times (TG/0.81) \times (1.52/HDL-C)$
Body roundness index	$364.2 - 365.5 \times \sqrt{1 - (WC/(2\pi))^2 / (0.5 \times height^2)}$
Waist-to-height ratio	WC/height
Conicity index	$WC/[0.109 \times \sqrt{(weight/height)}]$
Weight-adjusted waist index	WC/ $\sqrt{weight}$
A body shape index	$WC/[BMI^{2/3} \times height^{1/2}]$

BMI, body mass index; HDL-C, high-density lipoprotein cholesterol; TG, triglycerides; WC, waist circumference.

of metabolic abnormalities, it has been broadly used in previous studies to identify MONW. However, there is a possibility of missing early or subclinical metabolic risks in those with normal weight. Given that the TyG index is a surrogate marker of insulin resistance, it can complement MS criteria by identifying individuals with latent metabolic risk who may otherwise be overlooked. MS was confirmed on the basis of the criteria recommended in 2022 by the Korean Society for the Study of Obesity [19], which require the presence of three or more of the following components: (1) WC  $\geq 90$  cm in males and  $\geq 85$  cm in females; (2) systolic blood pressure  $\geq 130$  mmHg, diastolic blood pressure  $\geq 85$  mmHg, or use of antihypertensive medication; (3) fasting plasma triglyceride concentration  $>150$  mg/dL or the use of lipid-lowering medication; (4) fasting HDL-cholesterol concentration  $<40$  mg/dL in males and  $<50$  mg/dL in females, or the use of lipid-lowering medication; and (5) fasting plasma glucose concentration  $\geq 100$  mg/dL or the use of antihyperglycemic medication. Additionally, our previous study, which investigated the association between the TyG index and MONW in 4,721 older Korean adults, proposed sex-specific TyG index cutoff points of  $\geq 8.88$  for males and  $\geq 8.80$  for females [4]. These cutoff points were applied in the current study to identify MONW. The TyG index was calculated as  $\ln [\text{triglyceride concentration (mg/dL)} \times \text{fasting plasma glucose concentration (mg/dL)} / 2]$  [4,20].

### Statistical analysis

The independent t-test or the Mann-Whitney U-test was adopted to compare continuous variables between MONW and MHNW. For categorical variables, the chi-square test was applied. The results are presented as the means  $\pm$  standard deviations or numbers (percentages) (Table 2). Logistic regression was performed to assess the sex-specific odds ratios (ORs) concerning the relationships between body adiposity indices and MONW. The fully adjusted model accounted for several confounding factors, such as age, education level, household

income, smoking, drinking, physical activity (leisure-time or occupational physical activity, transfer time, and sedentary time), and nutrition (total, carbohydrate, protein, and fat caloric intake). The results are presented as ORs (95% confidence intervals). Statistical analyses were performed using SPSS software, ver. 26.0 (IBM, Inc.) (Table 3). To determine the optimal cutoff points for each sex-specific body adiposity index in identifying MONW individuals, we conducted receiver operating characteristic (ROC) curve analysis. The optimal cutoff points for each index were determined using the Youden index, which seeks to maximize the combined sensitivity and specificity. This analysis utilized MedCalc for Windows ver. 9.1.0.1 (MedCalc Corp.) (Table 4). For all analyses,  $P < 0.05$  was considered statistically significant.

## RESULTS

Table 2 presents the characteristics of the study participants. The mean age was  $70.0 \pm 6.4$  years for males and  $69.5 \pm 6.4$  years for females. In both sexes, all body adiposity indices in the MONW group were greater than those in their counterparts ( $P < 0.001$  for all). Similarly, all MS components and the TyG index were greater in the MONW group than in the MHNW group ( $P < 0.05$  for all).

Table 3 presents the sex-specific ORs for the associations between body adiposity indices and MONW. Among males, in the fully adjusted models, the VAI had ORs of 2.607 (2.168–3.135) for the middle tertile and 26.542 (21.158–33.294) for the highest tertile ( $P < 0.001$  for both). For WC, the ORs were 1.948 (1.645–2.307) and 4.538 (3.798–5.423) for the middle and highest tertiles, respectively ( $P < 0.001$  for both). The BRI yielded ORs of 1.981 (1.672–2.347) and 4.013 (3.357–4.797), whereas the WHtR presented nearly identical values, with ORs of 1.981 (1.672–2.347) and 4.016 (3.360–4.801) ( $P < 0.001$  for all). For the CI, the ORs were 1.853 (1.562–2.197) and 4.273 (3.554–5.137); for the WWI, they were 1.979 (1.668–2.348) and 3.847 (3.193–

**Table 2.** Characteristics of study subjects

Characteristic	Males				Females			
	All males (n=3,536)	MONW (n=1,760)	MHNW (n=1,776)	P-value	All females (n=4,310)	MONW (n=2,476)	MHNW (n=1,834)	P-value
VAI	4.036±5.209	5.712±6.911	2.374±1.079	<0.001 <sup>a</sup>	4.924±3.882	6.320±4.505	3.039±1.352	<0.001 <sup>a</sup>
WC (cm)	84.244±6.145	86.151±5.777	82.355±5.913	<0.001	79.897±6.106	81.480±5.878	77.759±5.748	<0.001
BRI	3.511±0.691	3.711±0.654	3.313±0.669	<0.001	3.792±0.813	3.987±0.797	3.529±0.758	<0.001
WHtR	0.507±0.037	0.517±0.034	0.496±0.036	<0.001 <sup>a</sup>	0.521±0.042	0.531±0.040	0.507±0.040	<0.001
CI	1.264±0.064	1.281±0.061	1.246±0.062	<0.001	1.248±0.072	1.265±0.070	1.226±0.067	<0.001
WWI	10.684±0.573	10.823±0.549	10.546±0.563	<0.001	10.989±0.683	11.134±0.669	10.793±0.653	<0.001
ABSI	0.0082±0.0004	0.0083±0.0004	0.0081±0.0004	<0.001	0.0081±0.0004	0.0082±0.0004	0.0080±0.0004	<0.001
BMI (kg/m <sup>2</sup> )	22.502±1.650	22.858±1.515	22.149±1.701	<0.001 <sup>a</sup>	22.470±1.620	22.737±1.545	22.108±1.649	<0.001 <sup>a</sup>
SBP (mmHg)	125.8±17.0	127.6±16.9	124.0±16.9	<0.001	127.4±17.5	130.2±17.0	123.6±17.5	<0.001
DBP (mmHg)	73.4±9.7	73.8±9.8	73.1±9.5	<0.05	73.3±9.2	73.7±9.3	72.7±9.1	<0.001
Glucose (mg/dL)	107.4±25.1	116.1±30.0	98.8±14.8	<0.001 <sup>a</sup>	103.5±23.6	109.8±28.0	95.0±11.2	<0.001 <sup>a</sup>
HDL-C (mg/dL)	48.6±12.5	45.5±12.2	51.8±12.0	<0.001	53.0±12.9	49.7±11.9	57.6±12.8	<0.001 <sup>a</sup>
Triglyceride (mg/dL)	130.4±99.8	172.4±125.2	88.8±28.3	<0.001 <sup>a</sup>	122.0±67.6	147.4±76.7	87.7±27.2	<0.001 <sup>a</sup>
TyG index	8.7±0.6	9.0±0.6	8.3±0.4	<0.001 <sup>a</sup>	8.6±0.5	8.9±0.5	8.3±0.3	<0.001 <sup>a</sup>
Age (yr)	70.0±6.4	69.8±6.3	70.3±6.4	<0.01	69.5±6.4	70.2±6.3	68.6±6.5	<0.001
Height (cm)	166.4±5.9	166.6±5.8	166.1±6.0	<0.01	153.5±5.8	153.6±5.7	153.4±6.0	0.213
Weight (kg)	62.4±6.5	63.6±6.3	61.2±6.5	<0.001	53.0±5.4	53.7±5.3	52.1±5.4	<0.001
Household income				0.220				<0.001
Low	1,181 (33.4)	592 (33.6)	589 (33.2)		1,738 (40.3)	1,052 (42.5)	686 (37.4)	
Lower-middle	1,064 (30.1)	552 (31.4)	512 (28.8)		1,180 (27.4)	688 (27.8)	492 (26.8)	
Upper-middle	709 (20.1)	343 (19.5)	366 (20.6)		785 (18.2)	425 (17.2)	360 (19.6)	
High	582 (16.5)	273 (15.5)	309 (17.4)		607 (14.1)	311 (12.6)	296 (16.1)	
Education level				0.468				<0.001
Primary school	1,198 (33.9)	602 (34.2)	596 (33.6)		2,417 (56.1)	1,467 (59.2)	950 (51.8)	
Middle school	646 (18.3)	309 (17.6)	337 (19.0)		684 (15.9)	389 (15.7)	295 (16.1)	
High school	1,003 (28.4)	515 (29.3)	488 (27.5)		812 (18.8)	430 (17.4)	382 (20.8)	
College	689 (19.5)	334 (19.0)	355 (20.0)		397 (9.2)	190 (7.7)	207 (11.3)	
Drinking				<0.001				<0.05
Never	1,042 (29.5)	461 (26.2)	581 (32.7)		2,550 (59.2)	1,508 (60.9)	1,042 (56.8)	
≤Once a week	1,336 (37.8)	661 (37.6)	675 (38.0)		1,549 (35.9)	854 (34.5)	695 (37.9)	
2–3 times/wk	614 (17.4)	339 (19.3)	275 (15.5)		143 (3.3)	74 (3.0)	69 (3.8)	
≥4 times/wk	544 (15.4)	299 (17.0)	245 (13.8)		68 (1.6)	40 (1.6)	28 (1.5)	
Smoking				<0.001				0.538
Never	709 (20.1)	291 (16.5)	418 (23.5)		4,059 (94.2)	2,325 (93.9)	1,734 (94.5)	
Former smoking	2,003 (56.6)	1,029 (58.5)	974 (54.8)		132 (3.1)	82 (3.3)	50 (2.7)	
Current smoking	824 (23.3)	440 (25.0)	384 (21.6)		119 (2.8)	69 (2.8)	50 (2.7)	
Medication								
Anti-hypertensive medication	1,402 (39.6)	919 (52.2)	483 (27.2)	<0.001	1,827 (42.4)	1,367 (55.2)	460 (25.1)	<0.001
Lipid-lowering medication	667 (18.9)	620 (35.2)	47 (2.6)	<0.001	1,342 (31.1)	1,186 (47.9)	156 (8.5)	<0.001
Anti-hyperglycemic medication	646 (18.3)	497 (28.2)	149 (8.4)	<0.001	723 (16.8)	636 (25.7)	87 (4.7)	<0.001
Physical activity								
LPA (min/wk)	62.0±158.9	58.8±156.6	65.2±161.1	0.237	31.3±98.5	23.4±80.6	42.0±117.7	<0.001 <sup>a</sup>
OPA (min/wk)	22.8±154.7	23.1±165.8	22.6±142.9	0.927	15.1±125.1	11.1±98.0	20.5±154.2	0.184
Transfer time (min/wk)	112.8±198.5	110.1±190.1	115.6±206.5	0.412	98.2±160.6	95.7±155.0	101.7±167.8	0.223
Sedentary time (min/wk)	3,220.6±1,521.3	3,298.7±1,508.2	3,143.2±1,530.7	<0.01	3,223.0±1,645.8	3,276.1±1,674.5	3,151.2±1,603.9	<0.05

(Continued on the next page)



**Table 2.** Continued

Characteristic	Males				Females			
	All males (n=3,536)	MONW (n=1,760)	MHNW (n=1,776)	P-value	All females (n=4,310)	MONW (n=2,476)	MHNW (n=1,834)	P-value
<b>Nutrition</b>								
Total energy intake (kcal/day)	1,978.2±742.8	1,934.8±710.3	2,021.1±771.5	<0.01	1,494.2±583.2	1,468.3±571.3	1,529.2±597.3	<0.01
Protein (g/day)	68.6±32.5	65.7±29.9	71.4±34.6	<0.001 <sup>a</sup>	51.6±25.6	50.3±24.6	53.4±26.9	<0.001 <sup>a</sup>
Fat (g/day)	36.4±26.8	35.0±24.9	37.8±28.5	<0.05 <sup>a</sup>	27.9±20.9	26.9±20.3	29.3±21.5	<0.001 <sup>a</sup>
Carbohydrate (g/day)	318.3±116.8	309.2±110.5	327.2±122.2	<0.001 <sup>a</sup>	257.1±103.3	254.0±101.7	261.2±105.3	<0.05

Values are presented as mean±standard deviation or number (%).

ABSI, a body shape index; BMI, body mass index; BRI, body roundness index; CI, conicity index; DBP, diastolic blood pressure; HDL-C, high-density lipoprotein cholesterol; LPA, leisure-time physical activity; MHNW, metabolically healthy and normal weight; MONW, metabolically obese but normal weight; OPA, occupational physical activity; SBP, systolic blood pressure; TyG index, triglyceride glucose index; VAI, visceral adiposity index; WC, waist circumference; WHtR, waist-to-height ratio; WWI, weight-adjusted-waist index.

<sup>a</sup>Mann-Whitney U-test was applied to assess the difference between groups.

**Table 3.** Sex-specific odds ratios for the association between body adiposity indices and MONW

Males (n=3,536)			Females (n=4,310)		
Variable	Unadjusted model	Fully adjusted model <sup>a</sup>	Variable	Unadjusted model	Fully adjusted model <sup>a</sup>
<b>VAI</b>			<b>VAI</b>		
LT (n=1,179)	Reference	Reference	LT (n=1,436)	Reference	Reference
MT (n=1,178)	2.600 (2.167–3.119)***	2.607 (2.168–3.135)***	MT (n=1,437)	2.329 (1.999–2.712)***	2.261 (1.938–2.638)***
HT (n=1,179)	25.951 (20.792–32.390)***	26.542 (21.158–33.294)***	HT (n=1,437)	21.868 (17.723–26.982)***	21.495 (17.373–26.594)***
<b>WC</b>			<b>WC</b>		
LT (n=1,182)	Reference	Reference	LT (n=1,437)	Reference	Reference
MT (n=1,195)	1.906 (2.770–4.301)***	1.948 (1.645–2.307)***	MT (n=1,440)	1.710 (1.475–1.982)***	1.683 (1.450–1.953)***
HT (n=1,159)	4.214 (3.546–5.008)***	4.538 (3.798–5.423)***	HT (n=1,433)	4.317 (3.679–5.064)***	4.060 (3.452–4.775)***
<b>BRI</b>			<b>BRI</b>		
LT (n=1,179)	Reference	Reference	LT (n=1,436)	Reference	Reference
MT (n=1,177)	1.966 (1.665–2.321)***	1.981 (1.672–2.347)***	MT (n=1,437)	1.851 (1.596–2.147)***	1.785 (1.536–2.076)***
HT (n=1,180)	3.617 (3.052–4.286)***	4.013 (3.357–4.797)***	HT (n=1,437)	3.638 (3.112–4.254)***	3.292 (2.793–3.879)***
<b>WHtR</b>			<b>WHtR</b>		
LT (n=1,178)	Reference	Reference	LT (n=1,436)	Reference	Reference
MT (n=1,179)	1.968 (1.667–2.324)***	1.981 (1.672–2.347)***	MT (n=1,437)	1.852 (1.597–2.148)***	1.786 (1.537–2.077)***
HT (n=1,179)	3.621 (3.055–4.292)***	4.016 (3.360–4.801)***	HT (n=1,437)	3.637 (3.110–4.253)***	3.290 (2.791–3.878)***
<b>CI</b>			<b>CI</b>		
LT (n=1,179)	Reference	Reference	LT (n=1,436)	Reference	Reference
MT (n=1,178)	1.758 (1.490–2.075)***	1.853 (1.562–2.197)***	MT (n=1,437)	1.784 (1.539–2.069)***	1.688 (1.453–1.962)***
HT (n=1,179)	3.667 (3.093–4.346)***	4.273 (3.554–5.137)***	HT (n=1,437)	3.734 (3.192–4.369)***	3.348 (2.839–3.948)***
<b>WWI</b>			<b>WWI</b>		
LT (n=1,178)	Reference	Reference	LT (n=1,436)	Reference	Reference
MT (n=1,179)	1.854 (1.571–2.187)***	1.979 (1.668–2.348)***	MT (n=1,437)	1.826 (1.575–2.117)***	1.731 (1.488–2.013)***
HT (n=1,179)	3.214 (2.716–3.084)***	3.847 (3.193–4.636)***	HT (n=1,437)	3.199 (2.741–3.734)***	2.845 (2.401–3.369)***
<b>ABSI</b>			<b>ABSI</b>		
LT (n=1,178)	Reference	Reference	LT (n=1,436)	Reference	Reference
MT (n=1,179)	1.727 (1.464–2.036)***	1.817 (1.535–2.154)***	MT (n=1,437)	1.593 (1.375–1.846)***	1.499 (1.290–1.742)***
HT (n=1,179)	3.097 (2.619–3.664)***	3.519 (2.930–4.226)***	HT (n=1,437)	3.293 (2.819–3.847)***	2.917 (2.474–3.440)***
<b>BMI</b>			<b>BMI</b>		
LT (n=1,178)	Reference	Reference	LT (n=1,436)	Reference	Reference
MT (n=1,179)	1.890 (1.603–2.228)***	1.983 (1.675–2.347)***	MT (n=1,437)	1.586 (1.368–1.837)***	1.622 (1.396–1.884)***
HT (n=1,179)	2.661 (2.253–3.143)***	2.838 (2.392–3.367)***	HT (n=1,437)	2.426 (2.085–2.824)***	2.464 (2.111–2.877)***

Values are presented as odds ratio (95% confidence interval).

ABSI, a body shape index; BMI, body mass index; BRI, body roundness index; CI, conicity index; HT, highest tertile; LT, lowest tertile; MONW, metabolically obese but normal weight; MT, middle tertile; VAI, visceral adiposity index; WC, waist circumference; WHtR, waist-to-height ratio; WWI, weight-adjusted-waist index.

<sup>a</sup>Fully adjusted model was adjusted for age, education level, house income, smoking, drinking, leisure-time or occupational physical activity, transfer time, sedentary time and total, carbohydrate, protein and fat intake calories.

\*P<0.05, \*\*P<0.01, \*\*\*P<0.001.

**Table 4.** Sex-specific AUC comparisons for body adiposity indices in identifying MONW

Variable	AUC	Cutoff	Sensitivity	Specificity	P-value
Males (n=3,536)					
VAI	0.826	>3.181	72.8	79.1	<0.001
WC	0.679	>84.300	64.6	61.7	<0.001
BRI	0.665	>3.433	68.0	57.2	<0.001
WHtR	0.665	>0.504	68.1	57.2	<0.001
CI	0.659	>1.280	51.9	70.7	<0.001
WWI	0.640	>10.547	69.5	50.5	<0.001
ABSI	0.635	>0.0829	51.5	69.0	<0.001
BMI	0.622	>22.221	69.9	48.6	<0.001
Females (n=4,310)					
VAI	0.802	>4.339	64.4	84.3	<0.001
WC	0.676	>80.400	57.5	67.4	<0.001
BRI	0.662	>3.675	64.2	59.4	<0.001
WHtR	0.662	>0.517	64.3	59.3	<0.001
CI	0.662	>1.257	55.9	68.6	<0.001
WWI	0.646	>10.931	62.2	59.9	<0.001
ABSI	0.645	>0.0822	48.2	73.4	<0.001
BMI	0.612	>22.576	59.7	57.6	<0.001

ABSI, a body shape index; AUC, area under the curve; BMI, body mass index; BRI, body roundness index; CI, conicity index; MONW, metabolically obese but normal weight; VAI, visceral adiposity index; WC, waist circumference; WHtR, waist-to-height ratio; WWI, weight-adjusted-waist index.

4.636); for the ABSI, they were 1.817 (1.535–2.154) and 3.519 (2.930–4.226); and for the BMI, they were 1.983 (1.675–2.347) and 2.838 (2.392–3.367) ( $P<0.001$  for all).

Among females, in the fully adjusted models, the VAI demonstrated ORs of 2.261 (1.938–2.638) for the middle tertile and 21.495 (17.373–26.594) for the highest tertile ( $P<0.001$  for both). For WC, the ORs were 1.683 (1.450–1.953) and 4.060 (3.452–4.775), respectively ( $P<0.001$  for both). The BRI had ORs of 1.785 (1.536–2.076) and 3.292 (2.793–3.879), whereas the WHtR yielded similar values: 1.786 (1.537–2.077) and 3.290 (2.791–3.878) ( $P<0.001$  for all). For the CI, the ORs were 1.688 (1.453–1.962) and 3.348 (2.839–3.948); for the WWI, they were 1.731 (1.488–2.013) and 2.845 (2.401–3.369); for the ABSI, they were 1.499 (1.290–1.742) and 2.917 (2.474–3.440); and for the BMI, they were 1.622 (1.396–1.884) and 2.464 (2.111–2.877) ( $P<0.001$  for all).

Table 4 presents the comparison of sex-specific ROC curves for body adiposity indices in identifying MONW. Among males, the VAI had the highest AUC of 0.826 (cutoff point: >3.181, sensitivity: 72.8, specificity: 79.1) ( $P<0.001$ ). WC had an AUC of 0.679 (cutoff point: >84.300, sensitivity: 64.6, and specificity: 61.7) ( $P<0.001$ ). Two indices, BRI and WHtR, demonstrated the same AUC of 0.665, with similar sensitivities (68.0 for BRI and 68.1 for WHtR) and the same specificity (57.2 for both) ( $P<0.001$  for both). The CI, WWI, and ABSI exhibited progressively lower AUCs in descending order, with varying sensitivities (51.9

for CI, 69.5 for WWI, and 51.5 for ABSI) and specificities (70.7 for CI, 50.5 for WWI, and 69.0 for ABSI) ( $P<0.001$  for all). BMI had the lowest AUC of 0.622 (cutoff point: >22.221, sensitivity: 69.9, specificity: 48.6) ( $P<0.001$ ).

Similarly, among females, the VAI had the highest AUC at 0.802 (cutoff point: >4.339, sensitivity: 64.4, specificity: 84.3) ( $P<0.001$ ). WC followed, with an AUC of 0.676 (cutoff point: >80.400, sensitivity: 57.5, specificity: 67.4) ( $P<0.001$ ). Three indices, BRI, WHtR, and CI, had the same AUC of 0.662, with varying sensitivities (64.2 for BRI, 64.3 for WHtR, and 55.9 for CI) and specificities (59.4 for BRI, 59.3 for WHtR, and 68.6 for CI) ( $P<0.001$  for all). The WWI and ABSI yielded similar AUCs of 0.646 and 0.645, respectively, with different sensitivities (62.2 for WWI and 48.2 for ABSI) and specificities (59.9 for WWI and 73.4 for ABSI) ( $P<0.001$  for both). BMI had the lowest AUC at 0.612 (cutoff point: >22.576, sensitivity: 59.7, specificity: 57.6) ( $P<0.001$ ).

## DISCUSSION

This cross-sectional study examined the associations between various WC-based body adiposity indices and MONW in older Korean adults and compared their discriminative ability. The major findings can be summarized as follows: (1) All WC-based body adiposity indices were associated with MONW in both sexes; however, only the VAI demonstrated a stronger as-

sociation than WC did. (2) Although all indices could identify MONW in both sexes, only VAI outperformed WC in terms of discriminative ability, whereas the other indices showed inferior performance. These findings suggest that the VAI may serve as the most effective indicator for identifying MONW among older Korean adults that are not underweight or obese.

Ruderman and colleagues [21] first introduced the concept of MONW, which is a subset of individuals with a BMI <25 kg/m<sup>2</sup> who exhibit hyperinsulinemia, insulin resistance, and other metabolic abnormalities typically observed in obese individuals. Later, in 1998, they further confirmed that MONW individuals are relatively common and face an elevated risk for type 2 diabetes and cardiovascular disease [22]. Following their foundational work, a growing body of research has demonstrated various health concerns associated with MONW in Asian older adults [23-25]. Despite such health risks associated with MONW, the need for a universally applicable identification index has long been overlooked in public health.

In the present study, the VAI had the strongest association and the highest discriminative ability for identifying MONW in both sexes. Prior to the present study, several previous studies have shown that the VAI is effective at detecting metabolic abnormalities. For example, a cross-sectional study by Motamed et al. [26] indicated that the VAI was more closely related to MS than other indices, such as WC, BMI, WHR, and WHtR, in both sexes. Compared with other indices, the VAI has also demonstrated a superior ability to discriminate MS [26]. Additionally, a longitudinal study by Kang et al. [27] revealed that individuals in the highest tertiles of the VAI presented a 2.26-fold greater risk for males and a 1.75-fold greater risk for females in terms of progression from a metabolically healthy status to a metabolically unhealthy status. The VAI also exhibited greater discriminative ability than WC did, regardless of sex [27].

Although direct comparisons between these previous studies and the current study are limited by differences in participant characteristics, such as ethnicity, age, and BMI ranges, the previous findings collectively support the utility of the VAI as a sex-independent indicator for identifying and predicting metabolic abnormalities [13]. This usefulness may be derived from the incorporation of both anthropometric (WC and BMI) and metabolic (triglyceride and HDL cholesterol) parameters in the VAI. This combination allows for a more comprehensive reflection of visceral adiposity accumulation and dysfunction than does WC alone [13]. However, it is important to note that the VAI includes triglyceride and HDL cholesterol, which are also

part of the operational definition of MONW. This overlap may lead to inflated observed associations and AUC values, limiting the interpretation of the VAI as a fully independent screening tool for metabolic risk.

In contrast, although the BRI, WHtR, CI, WWI, and ABSI were developed to improve assessments of abdominal obesity compared with WC or BMI, their associations with MONW and their discriminative abilities in older adults are relatively modest [12,14-16]. One possible explanation is that although these indices are helpful for general obesity screening, they may not adequately reflect the complex relationship between visceral fat and metabolic abnormalities in older adults. This is particularly relevant given that older individuals undergo physiological changes such as fat redistribution and loss of muscle mass, which complicate the obesity phenotype [28].

A key strength of this study is the extensive adjustment for demographic and lifestyle variables, enhancing the reliability of the associations observed. However, owing to the cross-sectional design, causality cannot be inferred. Longitudinal studies are needed to confirm these findings. Additionally, the inclusion criterion of BMI between 18.5 and 24.9 kg/m<sup>2</sup>, while consistent with standard definitions of normal weight, may not accurately reflect true adiposity or metabolic risk in older adults. Age-related physiological changes, such as sarcopenia and fat redistribution, are not captured by BMI alone and may have influenced the classification of MONW individuals in this population. Furthermore, since all participants were older Korean adults, it remains uncertain whether the results can be extrapolated to individuals of other ethnicities or in different countries. Measurements such as WC and WC-based indices used in this study are known to vary by ethnicity, age, and sex, highlighting the need for ethnicity-specific cutoff points. Finally, although MONW was defined using both MS criteria and the TyG index to improve sensitivity, their partial overlap, particularly concerning glucose and triglyceride components, may introduce redundancy. The TyG-based cutoff values, which were derived from a Korean sample, also require external validation before being applied more broadly.

In conclusion, our findings suggest that, compared with WC and other indices, the VAI may be the most effective WC-based index for identifying MONW in older adults that are not underweight or obese. Considering the limitations of BMI, WC, and other indices in capturing metabolic abnormalities, the VAI may serve as a more practical and accurate indicator for clinical screening and public health surveillance.

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## AUTHOR CONTRIBUTIONS

Dr. Bokun KIM had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis. The author reviewed this manuscript and agreed to individual contributions.

Conceptualization, data curation, formal analysis, methodology, validation, writing original draft, reviewing & editing: BK.

## CONFLICTS OF INTEREST

No existing or potential conflict of interest relevant to this article was reported.

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## DATA AVAILABILITY

Publicly available datasets were analyzed in this study. These data can be found here: The datasets generated during the current study are available in the 2014–2022 KHANES (<https://knhanes.kdca.go.kr/knhanes/main.do>) and analyzed during the current study are available from the corresponding author upon reasonable request.

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# The Effect of Resistance and Power Exercise Types on Maximum Strength and Muscular Endurance in University Students

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## ABSTRACT

**Background:** This study aimed to compare and analyze the effects of two training modalities—resistance training and power training—on maximal strength (one-repetition maximum, 1RM) and repetition-based muscular endurance in university students.

**Methods:** Fifty-seven university students (34 males and 23 females) participated in a 12-week training intervention. The program was divided into two phases: phase 1 (first 6 weeks) involved resistance training with a controlled 3-second concentric-eccentric tempo, while phase 2 (last 6 weeks) incorporated power training with 20-second maximal repetition sets. A total of 12 exercises were included, comprising eight 1RM-based and four repetition-based exercises. Measurements were conducted at three time points—T1 (pre), T2 (mid), and T3 (post)—and included assessments of body composition and exercise performance. Data were analyzed using repeated measures ANOVA and paired t-tests.

**Results:** Significant improvements ( $P < 0.001$ ) were observed across most 1RM-based exercises, particularly during the power training phase. Repetition-based exercises also showed positive trends, although performance declines were noted in the kettlebell exercise for some participants. In terms of body composition, skeletal muscle mass significantly increased across all participants, while male participants exhibited a slight decrease in body fat percentage.

**Conclusions:** These findings suggest that power training is highly effective in improving maximal strength within a short period and that the effectiveness of exercise interventions may differ depending on contraction velocity and training modality. Based on a practical training design for general university students, this study offers foundational insights for developing efficient physical training programs.

**Keywords:** Resistance training, Power training, Maximal strength, Muscular endurance, 1 Repetition maximal

## INTRODUCTION

In modern society, the issue of insufficient physical activity among university students is becoming increasingly serious [1]. This lack of activity has been associated with a heightened risk

of decreased physical fitness, obesity, and musculoskeletal disorders [2]. Particularly, many students experience a significant decline in physical activity upon entering university, which can lead directly to deteriorating health and a reduced quality of life in the long term [3]. Against this backdrop, the importance of

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strength training has been emphasized [4]. Resistance training enhances muscle hypertrophy and maximal strength [5-7], while power training, performed at high velocity with low resistance, improves neuromuscular efficiency and fast-twitch fiber recruitment [8,9]. As these modalities differ in stimulus, their effects on maximal strength also differ. Maximal strength, commonly measured via one-repetition maximum (1RM) [10-12], is a key indicator of performance. Therefore, this study aims to compare the effects of resistance and power training on university students' maximal strength.

A unique aspect of this study is the use of identical exercises across two phases. Phase 1 applied the "3-second rule" for controlled repetitions, while phase 2 focused on high-speed repetitions within 20 seconds. Muscle strength and body composition were measured pre-, mid-, and post-intervention to assess training effects.

## METHODS

This study involved 57 male and female students from Sogang University (mean age, 22.04 years). This study was approved by the Institutional Review Board at Sogang University (SGUIRB-A-1808-44) and all participants signed an informed consent form. Body composition (height, weight, muscle mass, %body fat) was measured using InBody (InBody Co.) at three time points (Table 1).

### Experimental design and procedure

A 12-week repeated-measures design was used, with 24 sessions held twice weekly. The first 6 weeks focused on resistance training with controlled concentric and eccentric phases ("3-second rule"), and the last 6 weeks on power training with explosive movements.

Participants were assessed at:

- T1 (pre): 1RM/repetition test+InBody
- T2 (mid): post-resistance training
- T3 (post): post-power training

### Exercise program structure

Twelve exercises were used:

- 1RM-based (8): leg press, leg curl, leg extension, shoulder press, chest press, fly, lat pull down, seated row
- Repetition-based (4): leg raise, sit-up, back extension, kettlebell

In phase 1, 1RM exercises followed the 3-second tempo, and repetitions were counted over 1 minute. Repetition-based exercises were done at a self-paced max effort for 1 minute. In phase 2, all exercises were performed as fast as possible for 20 seconds using the same load. Sessions included warm-up and cooldown, supervised by exercise specialists.

### Statistical analysis

SPSS ver. 25.0 (IBM Corp.) was used. Means, standard deviations, repeated measures ANOVA, and paired t-tests were applied. For statistical analysis, Friedman repeated measures ANOVA (a non-parametric test) and paired t-tests were conducted for each exercise type. Missing data were handled by listwise deletion. Significance was set at  $\alpha=0.05$ .

## RESULTS

### Analysis of one-repetition maximum-based exercises

All eight 1RM-based exercises showed statistically significant improvements across time ( $P<0.001$ ), with paired t-tests confirming gains from T1 to T3 (Fig. 1). This supports the effectiveness of power training on maximal strength. In contrast, some repetition-based exercises (leg raise, back extension, kettlebell) showed no significant improvement, possibly due to limited neuromuscular adaptation, core-focused nature, and inconsistent protocols between phases.

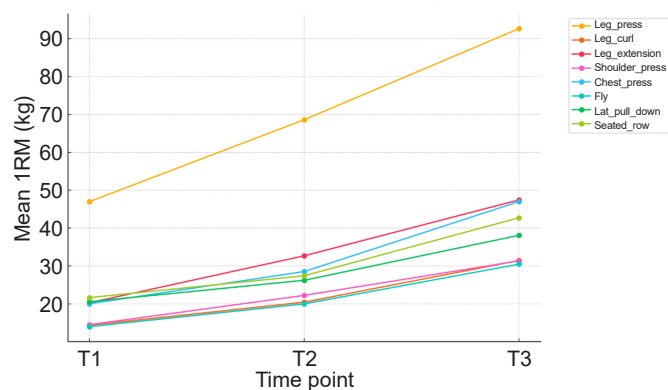
### Analysis of repetition-based exercises

Sit-up and leg raise improved significantly ( $P<0.05$ ) (Fig. 2). Back extension showed marginal significance, while kettlebell showed a Friedman effect but lacked paired t-test significance. Technical difficulty and execution variability likely influenced

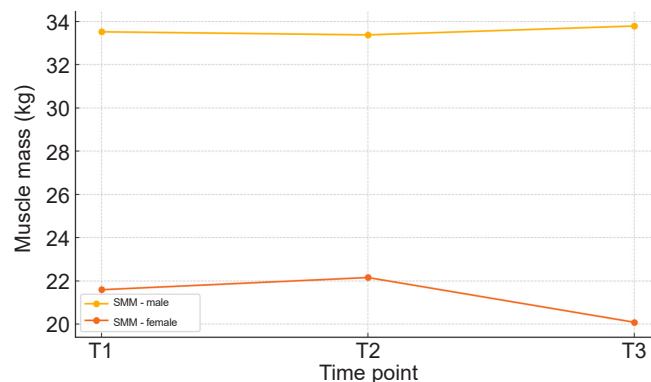
**Table 1.** Categorization of statistical methods used in the analysis

Statistical procedure	Included columns	Description
Mean and standard deviation	T1 mean, T3 mean	Summary of pre- and post-test mean values
Repeated measures ANOVA	Friedman $\chi^2$ , $P$ -value, N (repeated measures ANOVA)	Analysis of changes across three time points
Paired t-test	t-statistic, $P$ -value, N (paired t-test)	Statistical comparison between T1 and T3
Missing value handling (listwise)	N (repeated measures ANOVA), N (paired t-test)	Valid sample sizes after listwise deletion of missing values

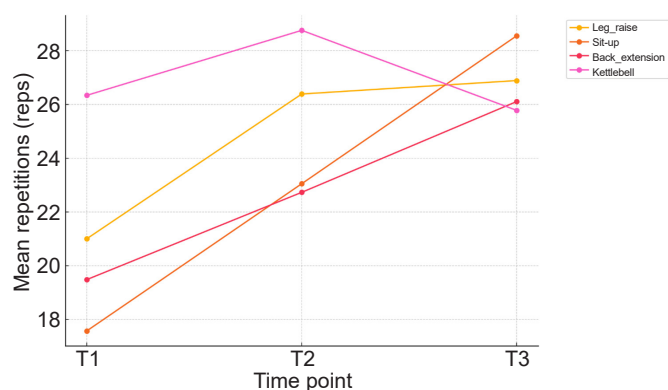
N, valid sample size; T1, pre-exercise; T3, post-exercise.



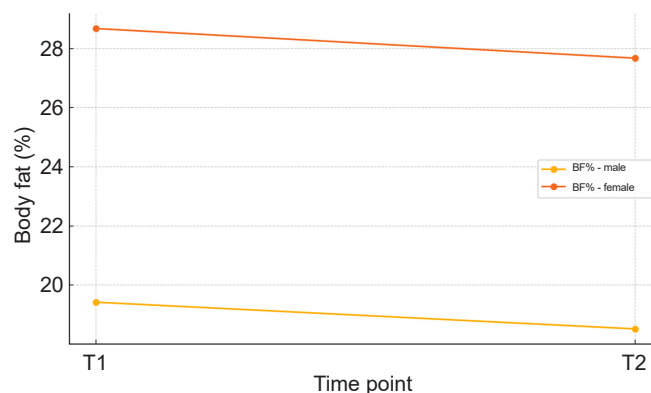
**Fig. 1.** Trends in mean one-repetition maximum (1RM) changes by exercise type. T1, pre-exercise; T2, mid-exercise; T3, post-exercise.



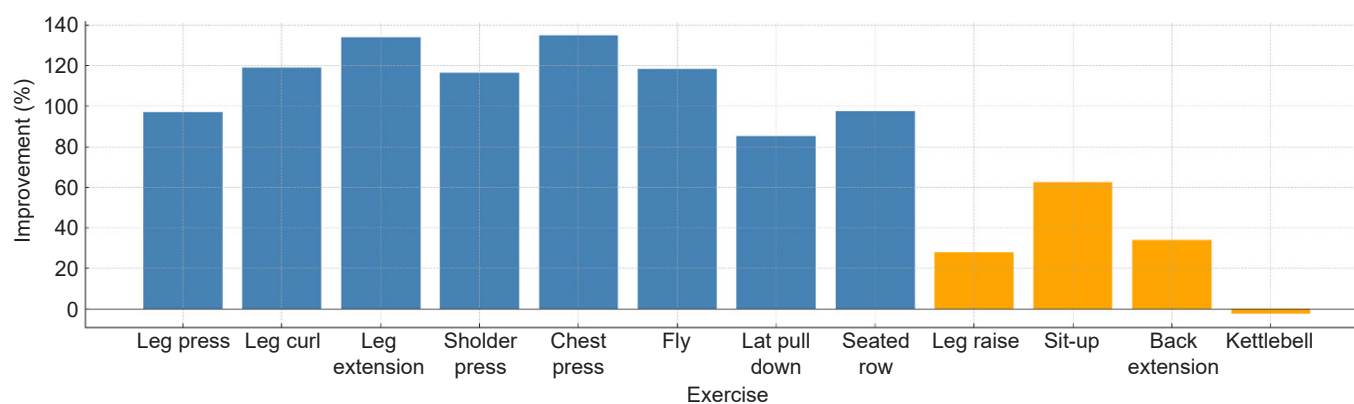
**Fig. 4.** Time-series changes in skeletal muscle mass (SMM) among male and female participants. T1, pre-exercise; T2, mid-exercise; T3, post-exercise.



**Fig. 2.** Changes in performance by exercise for repetition-based movements. T1, pre-exercise; T2, mid-exercise; T3, post-exercise.



**Fig. 5.** Time-series changes in body fat percentage (BF%) among male and female participants. T1, pre-exercise; T2, mid-exercise.



**Fig. 3.** Comparison of improvement rates by exercise from T1 (pre-exercise) to T3 (post-exercise).

these outcomes.

Improvement summary

Most 1RM exercises improved by over 100%, while endurance exercises had lower gains. Fig. 3 visualizes all exercise-specific improvements.

Body composition

Muscle mass increased significantly overall, with male participants showing a slight decrease in body fat (Fig. 4, 5). These trends are attributed to hormonal/metabolic responses to training.

Statistical overview

Table 2 summarizes key results for all exercises. The leg extension showed high significance ( $\chi^2=49.79$ ,  $t(30)=-8.97$ ,  $P<0.001$ ). Non-parametric methods were used when normality was violated. Training phases were clearly differentiated: the first phase applied time-under-tension (3-second rule), and the second emphasized explosive 20-second bouts. Performance differences, especially in kettlebell work, suggest the importance of both neuromuscular stimulus and technical proficiency.

DISCUSSION

This study investigated the effects of resistance and power training on maximal strength (1RM), endurance performance (repetition-based), and body composition (skeletal muscle mass and body fat percentage) in a population of university students. Among the 12 exercises evaluated, 11 showed significant

improvements, with particularly notable performance gains observed in exercises such as leg extension, chest press, and sit-up. The average improvement rate for the 1RM-based exercises exceeded 100%, indicating a pronounced effect of the training intervention over a relatively short period.

Resistance training has been shown to enhance muscle hypertrophy, neuromuscular efficiency, and muscle fiber cross-sectional area [13]. In contrast, power training emphasizes explosive speed and fast-twitch fiber recruitment, aiming to maximize repetitions within a short time window. In this study, both forms of training contributed to improved strength outcomes. Given that resistance and power training were applied sequentially rather than as isolated sessions, the observed effects may be interpreted as resulting from a combined or hybrid training model.

Skeletal muscle mass changes differed by sex; male participants generally showed consistent increases, whereas female participants displayed maintenance or slight decreases after the second measurement. This pattern is likely related to hormonal differences [14] and muscle fiber distribution characteristics, aligning with previous findings that men exhibit more rapid muscular adaptations to high-load training. Body fat percentage gradually decreased in both sexes, presumably due to increased energy expenditure and enhanced metabolic efficiency.

Not all exercises showed improvement, however. The kettlebell exercise was the only activity that exhibited a decline in average repetitions, a result that was not statistically significant. This outcome may be attributed to the technical difficulty of the movement, insufficient individual proficiency, or inconsistencies in exercise intensity prescription [15]. Accumulated fatigue

Table 2. Statistical analysis results by exercise type

Exercise	Friedman $\chi^2$ (repeated measures ANOVA)	P-value (repeated measures ANOVA)	N (repeated measures ANOVA)	T1 mean	T3 mean	t-statistic (paired t-test)	P-value (paired t-test)	N (paired t-test)
Leg press	44.486	<0.001	30	43.55	88.84	-7.114	<0.001	31
Leg curl	38.71	<0.001	31	13.23	27.19	-6.673	<0.001	31
Leg extension	49.791	<0.001	31	19.35	46.94	-8.972	<0.001	31
Shoulder press	43.521	<0.001	30	14.03	29.32	-5.84	<0.001	31
Chest press	50.643	<0.001	30	18.55	44.35	-9.628	<0.001	31
Fly	43.146	<0.001	30	13.39	27.48	-6.093	<0.001	31
Lat pull down	38.242	<0.001	29	19.5	34.8	-6.706	<0.001	30
Seated row	35.958	<0.001	28	20.52	38.1	-7.433	<0.001	29
Leg raise	8.237	0.0163	15	19.5	25.75	-2.188	0.0449	16
Sit-up	17.238	<0.001	16	14.56	27.69	-4.132	<0.001	16
Back extension	8.842	0.012	15	18.88	26.0	-3.148	0.0066	16
Kettlebell	11.828	0.0027	15	28.93	26.33	1.033	0.319	15

N, valid sample size; T1, pre-exercise; T3, post-exercise.

and reduced movement accuracy may also have influenced the decrease in performance [16]. These findings highlight the need to reassess exercise-specific responsiveness and ensure clearly defined quantification criteria during program design.

Furthermore, among the repetition-based exercises, both leg raise and back extension did not exhibit statistically significant changes. This may be attributed to the inherent difficulty in providing a uniform training stimulus across participants. Specifically, the leg raise exercise involves lower abdominal activation without upper body stabilization, making performance highly dependent on individual core strength levels [17]. In the case of the back extension, variations in range of motion and the degree of bodyweight support across individuals hinder precise quantification of repetitions [18]. These factors likely contributed to the lack of statistical significance, suggesting the need for more refined intensity control methods for such exercises in future studies.

The results of this study show that resistance exercise effectively enhances muscular strength [19]. Furthermore, incorporating aerobic components into the training program appears to promote broader functional improvements [20]. This study provides empirical evidence that short-term, high-intensity interventions can significantly enhance maximal strength even in healthy college populations.

Nonetheless, several limitations should be acknowledged. First, the use of voluntary participants may limit the generalizability of the findings. Second, the independent effects of resistance and power training could not be fully isolated due to the sequential design of the intervention. Third, potential confounding variables such as prior training experience and physical activity levels outside the study were not controlled. Additionally, missing data were present for certain exercises and were addressed using listwise deletion.

Despite these limitations, the study presents a practical exercise program model aimed at improving maximal strength and endurance in a general college student population. The circuit-style intervention—incorporating a mix of strength- and repetition-based exercises—is suitable even for beginners and can be adapted for personalized programming based on sex and body composition.

A key strength of the study lies in its unique experimental design, which applied different training modalities to the same set of exercises. In the first phase, resistance training was conducted using the “3-second rule,” with each repetition performed over 1 minute at a controlled tempo (1.5 seconds concentric+1.5 sec-

onds eccentric). In the second phase, the same exercises were performed for 20 seconds using a power training format aimed at maximal repetition. By distinguishing between 1RM-measurable exercises (8 types) and repetition-based exercises (4 types), the study was able to comprehensively evaluate both strength and muscular endurance. This approach enabled a clear investigation into the combined effects of movement speed, duration, and repetition count on muscle development. It also provided a basis for examining the interaction between maximal strength and muscular endurance. Such a distinction is expected to serve as a meaningful framework not only for understanding improvements in muscular strength, but also for identifying neuromuscular and metabolic responses that differ according to training modality. Ultimately, the experimental framework suggests that a hybrid training program—combining resistance and power components—may yield superior outcomes compared to using either approach in isolation.

## Practical implications and recommendations

Based on the findings of this study, the following practical recommendations are proposed:

- 1) A combined circuit training program incorporating both resistance and power exercises is effective for improving physical fitness in university students.
- 2) Individualization of training variables such as intensity and repetitions is essential, taking into account sex and baseline fitness levels.
- 3) Technically demanding exercises (e.g., kettlebell movements) should be accompanied by proper instruction and repeated practice to ensure safety and efficacy.
- 4) Future studies should control for external factors such as training cycle, intensity modulation, dietary habits, and psychological motivation.
- 5) Longitudinal research is needed to assess the sustainability and long-term maintenance of training effects.

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## AUTHOR CONTRIBUTIONS

Dr. Daihyuk CHOI had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis. All authors reviewed this manuscript and agreed to individual contributions.

Conceptualization: JWK and DC. Data curation: JWK, HJ, and KAC. Formal analysis: JWK. Investigation: JWK. Methodology: JWK. Project Administration: JWK and DC. Resources: DC, HJ, and KAC. Supervision: DC. Validation: JWK, HJ, and KAC. Visualization: JWK. Writing—original draft: JWK. Writing—review & editing: DC, HJ, and KAC.

## CONFLICTS OF INTEREST

No existing or potential conflict of interest relevant to this article was reported.

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None.

## DATA AVAILABILITY

The data presented in this study are available upon reasonable request from the corresponding author.

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# Examining the Mediating Effect of Resilience on Depression and Suicidal Ideation in Older Adults Living Alone

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## ABSTRACT

**Background:** This study explores whether resilience mediates the effect of depression on suicidal ideation among older adults living alone. It aims to clarify the role of resilience as a psychological protective factor and provide foundational data for developing suicide prevention strategies tailored to this vulnerable population.

**Methods:** Eighty-four older adults aged 70–80 years, living alone and receiving specialized senior care services in three Korean localities, participated in the study. Depression was assessed using the Geriatric Depression Scale, suicidal ideation with the Scale for Suicidal Ideation, and resilience using the resilience subscale of the Psychological Capital Questionnaire. SPSS 28.0 was used to conduct descriptive statistics, correlation analysis, multiple regression analyses following the Baron and Kenny causal-steps approach, and the Sobel test to assess mediation effects.

**Results:** Depression was positively correlated with suicidal ideation ( $r=0.287$ ,  $P<0.01$ ) and negatively correlated with resilience ( $r=-0.315$ ,  $P<0.01$ ). Suicidal ideation also showed a significant negative correlation with resilience ( $r=-0.359$ ,  $P<0.01$ ). Multiple regression analyses following the Baron and Kenny causal-steps approach confirmed that resilience partially mediated the relationship between depression and suicidal ideation. The Sobel test supported this result ( $Z=2.042$ ,  $P<0.05$ ), indicating that resilience functions as a psychological buffer.

**Conclusions:** This study confirms the importance of resilience in reducing suicidal ideation among older adults living alone. The findings highlight the need for effective, multidimensional intervention strategies that enhance resilience, not only through psychological approaches but also through broader policy and environmental support.

**Keywords:** Psychological resilience, Depression, Suicide prevention, Protective factors, Aged

## INTRODUCTION

South Korea has rapidly advanced to become the world's 13th-largest economy; however, it faces a serious social issue as the country with the highest suicide rate among Organisation

for Economic Co-operation and Development (OECD) nations. According to Statistics Korea [1], the suicide rate in 2023 was 27.3 per 100,000 population, an increase of 2.2 from the previous year. Suicide remains the leading cause of death by external factors across all age groups aged 10 and above, highlighting the

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severity of the problem. Notably, the suicide rates are highest among those aged 80 and older (59.4 per 100,000) and those in their 70s (39.0 per 100,000), underscoring the growing suicide issue in the elderly population.

Among them, older adults living alone are considered a particularly vulnerable group to depression and suicide risk due to complex factors such as social isolation, chronic illness, and loneliness. According to the 2023 national survey on older adults [2] conducted by the Ministry of Health and Welfare, 2% of older adults living alone reported having experienced suicidal thoughts, which is four times higher than that of older adults living with a spouse (0.5%) and approximately 3.3 times higher than those living with children (0.6%). Loneliness accounted for the highest proportion of reasons for suicidal ideation (28.3%), and the prevalence of depressive symptoms (16.1%) was more than twice as high as that of older adults living with a spouse (7.8%). These findings indicate that older adults living alone constitute a high-risk group for suicide, drawing attention to the relationship between depression and suicidal ideation in this population.

However, not all experiences of depression lead to suicidal ideation. In the same survey, while 16.1% of older adults living alone reported experiencing depression, only 2.0% reported having suicidal thoughts. This suggests the presence of psychological protective factors that prevent the progression from depression to suicidal ideation, even when individuals are exposed to various stressors. Identifying such protective factors provides a crucial theoretical foundation for developing practical intervention strategies for suicide prevention.

Recently, resilience has attracted attention as one of the key psychological protective factors. Resilience refers to an individual's ability to respond flexibly and recover psychologically when faced with stress or adversity, potentially buffering the relationship between depression and suicide. Depression has been identified as a major predictor of suicidal ideation [3], while resilience has been reported as a protective factor that reduces suicidal thoughts [4]. Kwon [5] found a moderating effect of resilience in the relationship between depression and suicidal intent, and Kim [6] also reported a moderating effect of resilience in the relationship between loneliness and suicidal ideation among older adults living alone. These findings suggest that resilience may alter the intensity of suicidal responses under stressful circumstances.

Additionally, according to Fredrickson's broaden-and-build theory of positive emotions, positive emotions such as joy, in-

terest, contentment, and love promote cognitive flexibility and resilience. Increased resilience, in turn, activates positive emotions, facilitating psychological recovery and adaptation [7-9]. From this perspective, resilience may function as a protective factor in the relationship between depression and suicidal ideation.

These recent research trends indicate an expansion in the understanding of suicide from a risk-factor-centered approach to one that emphasizes protective factors. However, in South Korea, studies focusing on protective factors among high-risk groups such as older adults living alone remain limited.

Meanwhile, previous studies on resilience have primarily reported its buffering effect, whereby resilience moderates the intensity of the response between depression and suicidal ideation. However, there is a lack of domestic research examining the mediating role of resilience in the relationship between depression and suicidal ideation among older adults living alone. Therefore, this study focuses on the possibility that resilience functions not merely as a moderator of effect strength but as an internal psychological mechanism operating in the process through which depression leads to suicidal ideation.

Accordingly, this study aims to gain a deeper understanding of the internal psychological mechanism of resilience by exploring whether depression influences resilience, which in turn acts as a mediating factor affecting suicidal ideation. Through this, the study seeks to elucidate the psychological mechanism of resilience influencing suicidal ideation among older adults living alone and to provide foundational data for the development of effective suicide prevention interventions based on resilience.

## METHODS

### Study design

This study employed a cross-sectional survey design to examine the impact of depression on suicidal ideation among older adults living alone and to explore the mediating role of resilience in this relationship.

### Participants

The participants in this study were adults in their 70s and 80s who were living alone and registered with Senior Customized Care Service institutions in three local municipalities in South Korea. Based on statistical evidence that suicide rates are highest among individuals in their 70s and 80s across all age groups [1], this study conducted a survey targeting older adults within

this age range. All participants were fully informed about the purpose of the study and voluntarily provided written consent to participate. The required sample size was calculated using G\*Power 3.1 software to ensure the adequacy for hierarchical regression analysis. The parameters were set as an effect size of 0.15, a significance level ( $\alpha$ ) of 0.05, and statistical power ( $1-\beta$ ) of 0.80, resulting in a minimum required sample size of 68. A total of 90 questionnaires were collected; after excluding six responses due to low response reliability, data from 84 participants were included in the final analysis.

## Measures

### Suicidal ideation

Suicidal ideation was assessed using the Scale for Suicidal Ideation developed by Beck et al. [10] and adapted into Korean by Shin et al. [11]. The scale comprises 19 items, each rated on a 3-point Likert scale. Total scores range from 0 to 38, with higher scores indicating greater suicidal ideation. Among adults, a score of 9 or above is generally interpreted as reflecting a heightened level of suicidal ideation compared to age-matched norms. The Scale for Suicidal Ideation was reported as Cronbach's  $\alpha=0.87$  in Shin et al.'s study [11] and 0.91 in the present study.

### Depression

Depression was assessed using the Geriatric Depression Scale developed by Sheikh and Yesavage [12] and adapted into Korean by Park et al. [13]. The scale comprises of 15 items with dichotomous responses coded as "yes" (1 point) or "no" (0 points). Total scores range from 0 to 15, with higher scores reflecting greater depressive symptoms. A score of 6 or above is generally interpreted as indicative of clinically relevant depression. The Geriatric Depression Scale was reported as Cronbach's  $\alpha=0.94$  in Sheikh and Yesavage's study [12] and 0.61 in the present study.

### Resilience

Resilience was assessed using the resilience subscale (5 items) of the Psychological Capital Questionnaire originally developed by Luthans et al. [14] and subsequently adapted and revised by Kwon [5]. The original scale questionnaire comprises four subscales: self-efficacy, hope, optimism, and resilience; however, this study employed only the resilience subscale in accordance with the research objectives. Items were rated on a 5-point

Likert scale 1 (strongly disagree) to 5 (strongly agree), with total scores ranging from 5 to 25. Higher scores indicate a greater level of resilience. The reliability of the resilience subscale was reported as Cronbach's  $\alpha=0.91$  in Kwon's study [5] and 0.70 in the present study.

### General characteristics

This study included five general characteristic variables—sex, age, educational level, health status, and income level—that have been reported in previous studies to influence suicidal behavior among older adults living alone [2].

### Data collection

Data was collected between January 2 and January 31, 2025, from 84 older adults aged 70 to 80 years who live alone and voluntarily consented to participate in the study. Participants were recruited from specialized service providers affiliated with three local municipalities in South Korea.

### Data analysis

Data was analyzed using IBM SPSS Statistics for Windows, Version 28.0. The following statistical procedures were employed:

- 1) Descriptive statistics (means and standard deviations) were calculated to assess levels of depression, suicidal ideation, and resilience among the participants.
- 2) Independent t-tests and one-way ANOVA were used to examine differences in key variables according to general characteristics.
- 3) Pearson correlation analysis was performed to explore the relationships among the main variables.
- 4) Multiple regression analyses following the Baron and Kenny causal-steps approach analysis and the Sobel test were performed to verify the mediating effect of resilience in the relationship between depression and suicidal ideation.

### Ethical consideration

Prior to data collection, ethical approval for this study was obtained from the Institutional Review Board (IRB) of Gachon University, with which the researcher is affiliated (IRB Approval No.: 1044396-202410-HR-175-01). All research procedures were conducted in accordance with established ethical standards for human subjects research.

## RESULTS

### General characteristics of participants

The general characteristics of the 84 participants in this study are as follows. The majority were female, accounting for 82.1% ( $n=69$ ), while males comprised 17.9% ( $n=15$ ). In terms of age, 64.3% ( $n=54$ ) were in their 70s, and 35.7% ( $n=30$ ) were in their 80s, indicating a higher proportion of participants in their 70s. Regarding educational background, the largest groups were those with no formal education and those who had completed elementary school, each accounting for 32.1% ( $n=27$ ), suggesting that the participants generally had low educational attainment. Regarding perceived health status, 69.0% ( $n=58$ ) reported their health as “poor.” For monthly income, the most frequently reported range was between 400,000 and 800,000 KRW, cited by 66.7% ( $n=56$ ) of participants (Table 1).

### Differences in depression and suicidal ideation according to general characteristics

An analysis of depression scores by participants’ sociodemographic characteristics revealed no statistically significant differences by sex ( $t=0.349$ ,  $P=0.728$ ), age group ( $t=0.746$ ,  $P=0.458$ ), education level ( $F=1.305$ ,  $P=0.275$ ), perceived health status ( $F=2.205$ ,  $P=0.117$ ), or monthly income ( $F=1.928$ ,  $P=0.152$ ) (Table 1).

Similarly, suicidal ideation did not differ significantly by sex ( $t=1.759$ ,  $P=0.082$ ), age group ( $F=1.107$ ,  $P=0.296$ ), education level ( $F=1.018$ ,  $P=0.403$ ), perceived health status ( $F=0.201$ ,

$P=0.818$ ), or monthly income ( $F=1.747$ ,  $P=0.181$ ). These findings suggest that the sociodemographic variables examined in this study were not significantly associated with differences in depression or suicidal ideation among older adults living alone (Table 1).

### Levels and correlations among depression, suicidal ideation, and resilience

The mean scores for the main variables were as follows: depression,  $9.33\pm2.44$ ; suicidal ideation,  $13.06\pm7.73$ ; and resilience,  $15.52\pm3.88$  (Table 2). Pearson correlation analysis revealed a significant positive correlation between depression and suicidal ideation ( $r=0.287$ ,  $P<0.01$ ), as well as significant negative correlations between depression and resilience ( $r=-0.315$ ,  $P<0.01$ ) and between suicidal ideation and resilience ( $r=-0.359$ ,  $P<0.01$ ) (Table 2). These results indicate that higher levels of depression are associated with increased suicidal ideation and lower resilience, while higher resilience is associated with lower suicidal ideation.

**Table 2.** Means, SDs, and correlations of key variables ( $n=84$ )

Variable	Mean $\pm$ SD	Depression	Suicidal ideation	Resilience
Depression	9.33 $\pm$ 2.44	1		
Suicidal ideation	13.06 $\pm$ 7.73	0.287**	1	
Resilience	15.52 $\pm$ 3.88	-0.315**	-0.359**	1

SD, standard deviation.

\*\* $P<0.01$ .

**Table 1.** General characteristics of participants ( $n=84$ )

Characteristic	Category	n (%)	Depression		Suicidal ideation	
			t	P	t	P
Sex	Male	15 (17.9)	0.349	0.728	1.759	0.082
	Female	69 (82.1)				
Age (yr)	70s group	54 (64.3)	0.746	0.458	1.107	0.296
	80s group	30 (35.7)				
Education level	No formal education (illiterate)	27 (32.1)	1.305	0.275	1.018	0.403
	Elementary school graduate	27 (32.1)				
	Middle school graduate	12 (14.3)				
	High school graduate	14 (16.7)				
	College graduate or higher	4 (4.8)				
Perceived health status	Poor health	58 (69.0)	2.205	0.117	0.201	0.818
	Fair health	17 (20.2)				
	Good health	9 (10.7)				
Income (won)	Under 400,000	11 (13.1)	1.928	0.152	1.747	0.181
	From 400,000 to 800,000	56 (66.7)				
	80,000 or more	17 (20.2)				
Total		84 (100)				

The sum of the percentages does not equal 100% because of rounding.

### Mediating effect of resilience on the relationship between depression and suicidal ideation

To examine the mediating effect of resilience on the relationship between depression and suicidal ideation among older adults, this study employed the hierarchical regression approach proposed by Baron and Kenny [15]. The statistical significance of the mediating effect was further verified using the Sobel test.

The mediation analysis followed Baron and Kenny's three-step procedure. In the first step, the effect of the independent variable (depression) on the mediating variable (resilience) was tested. In the second step, the effect of depression on the dependent variable (suicidal ideation) was examined. In the third step, both depression and resilience were simultaneously entered into the regression model to assess their respective effects on suicidal ideation, thereby determining whether resilience mediated the relationship between depression and suicidal ideation.

In the first step of the mediation analysis, depression had a significant negative effect on resilience ( $t=-3.005$ ,  $P<0.01$ ), accounting for 8.8% of the variance in resilience. The model was statistically significant ( $F(1, 82)=9.031$ ,  $P<0.01$ ). In the second step, depression exhibited a significant positive effect on suicidal ideation ( $t=2.713$ ,  $P<0.01$ ), explaining 7.1% of the variance. This model was also statistically significant ( $F(1, 82)=7.358$ ,  $P<0.01$ ). In the third step, when both depression and resilience were entered simultaneously into the regression model, resilience demonstrated a significant negative effect on suicidal ideation ( $t=-2.781$ ,  $P<0.01$ ), while the effect of depression on suicidal ideation was reduced to marginal significance ( $t=1.802$ ,  $P<0.05$ ). The Sobel test further confirmed the significance of the mediating effect, yielding a Z-value of 2.042 ( $P<0.05$ ). These findings indicate that resilience partially mediates the relationship between depression and suicidal ideation (Table 3, Fig. 1).

## DISCUSSION

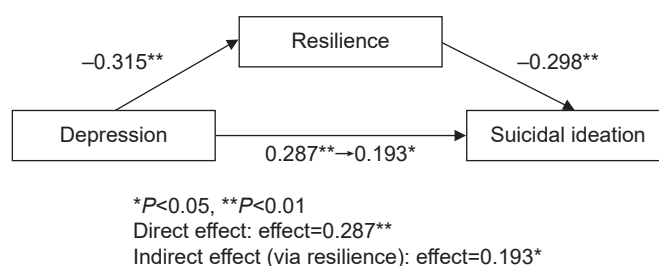
The main purpose of this study was to explore whether resilience mediates the relationship between depression and suicidal

ideation among older adults living alone. The main findings are discussed as follows.

First, an analysis of participants' demographic characteristics revealed that the majority were female (82.1%,  $n=69$ ), with the most represented age group being those in their 70s (64.3%). Regarding educational attainment, illiteracy and elementary school education were the most prevalent, each comprising 32.1% of the sample, followed by middle school, high school, and college graduates in descending order.

In terms of perceived health status, 69.0% of participants rated their health as "poor," and the most frequently reported monthly income was between 400,000 and 800,000 KRW (66.7%). These findings suggest that the older adults living alone in this study were characterized by low levels of education, poor perceived health, and financial hardship. Moreover, the average scores for depression and suicidal ideation were  $9.33\pm2.44$  and  $13.06\pm7.73$ , respectively, both exceeding the clinical cutoff points for risk ( $\geq 6$  for depression and  $\geq 9$  for suicidal ideation). These results indicate that this population is at elevated risk for both depression and suicidal ideation, potentially influenced by intersecting factors such as limited educational attainment, health vulnerability, and economic deprivation—findings that align with prior research [2].

Second, the analysis of differences in depression and suicidal ideation by demographic characteristics revealed no statistically significant differences across any of the examined variables,



**Fig. 1.** Mediating effect of resilience on depression and suicidal ideation in older adults living alone.

**Table 3.** Mediating effect of resilience on the influence of depression on suicidal ideation

	Independent variable	Dependent variable	B	SE	$\beta$	t	F (P)	R <sup>2</sup>	Adj. R <sup>2</sup>
Step 1	Depression	Resilience	-0.503	0.167	-0.315	-3.005	9.031**	0.099	0.088
Step 2	Depression	Suicidal ideation	0.910	0.335	0.287	2.713	7.358**	0.082	0.071
Step 3	Depression	Suicidal ideation	0.612	0.340	0.193	1.802	7.847**	0.016	0.014
	Resilience	Suicidal ideation	-0.592	0.213	-0.298	-2.781			

Adj., adjusted; SE, standard error.

\*\* $P<0.01$ , Sobel test statistic: 2.042,  $P<0.05$ .



including sex, age group, educational level, perceived health status, and income. This finding suggests that, among older adults living alone aged 70 and above, the levels of depression and suicidal ideation may be more strongly influenced by psychological rather than demographic factors [16]. These findings imply that while demographic background should be considered in interventions addressing depression and suicidal ideation among elderly living alone, interventions focusing primarily on intrapersonal psychological resources—such as resilience—and psychosocial factors may be more effective.

Third, correlation analysis among the key variables demonstrated that depression was significantly positively correlated with suicidal ideation ( $r=0.287$ ,  $P<0.01$ ) and significantly negatively correlated with resilience ( $r=-0.315$ ,  $P<0.01$ ). These results indicate that higher levels of depression are associated with increased suicidal ideation and lower levels of resilience. Furthermore, suicidal ideation was significantly negatively correlated with resilience ( $r=-0.359$ ,  $P<0.01$ ), suggesting that individuals with greater resilience tend to report lower levels of suicidal ideation.

These findings support the conceptualization of resilience as a protective psychological resource that mitigates the adverse effects of depression and suicidal ideation among older adults living alone. They underscore the potential value of resilience-enhancing interventions in suicide prevention efforts within this population.

Lastly, the analysis of the mediating effect of resilience on the relationship between depression and suicidal ideation among older adults living alone confirmed that resilience partially mediates this relationship. This finding suggests that depression influences suicidal ideation indirectly through the psychological resource of resilience, rather than exerting a direct effect. These findings are consistent with previous studies [17] that reported a negative association between resilience and suicidal ideation, as well as evidence indicating that ego-resilience significantly impacts suicidal ideation [4,6,18].

On the other hand, Kwon [5] reported that resilience functions as a moderator in the relationship between depression and suicidal ideation among older adults, which differs functionally from the mediating effect identified in this study. This discrepancy can be understood by distinguishing between buffering effects and indirect effects, which represent distinct functional pathways. Nevertheless, both perspectives underscore resilience as a psychological protective factor. In this study, resilience was identified as a mediating mechanism that attenuates the

pathway through which depression leads to suicidal ideation. This finding contrasts with the pathway in which hopelessness elevates suicide risk [16] and further underscores resilience as a critical psychological resource for suicide prevention. Relatedly, Cross [19] defined resilience as the emotional recovery capacity that protects against suicide risk by enabling individuals to withstand and recover from stressful situations. Furthermore, Cha and Lee [18] reported that problem-solving ability, a component of resilience, plays a role in reducing suicidal ideation. These findings reaffirm that resilience is a key psychological resource that alleviates suicide risk.

Meanwhile, recent studies have expanded the concept of resilience beyond individual internal traits to a more multidimensional, system-wide capacity for recovery. Peeters et al. [20] redefined resilience as the ability of a system to recover and adapt after exposure to stressors and proposed understanding it from a multilevel (systemic) perspective. Specifically, to support resilience in older adults with cognitive decline, they emphasized the need for integrated strategies at the macro level, including organizational, policy, and community resources, going beyond psychological and cognitive interventions. This perspective suggests that interventions aimed at enhancing resilience should extend beyond simple individual-centered approaches to more practical and feasible multidimensional strategies.

Therefore, given that resilience was identified as a significant mediating variable in the relationship between depression and suicidal ideation among older adults aged 70 and above living alone, the development of integrated intervention strategies at both micro and macro levels is warranted to enhance resilience in this population. For physically and economically vulnerable older adults living alone, individual psychological counseling or cognitive training alone may be insufficient to effectively strengthen resilience. Accordingly, a multidimensional, community-based approach should be implemented in parallel to address the complex and interrelated factors contributing to suicide risk.

Based on the above discussion, the following recommendations are proposed. First, resilience has been identified as a key psychological protective factor that buffers the effects of depression and suicidal ideation among older adults aged 70 and above who are at risk of dying alone. Future research should focus on developing and evaluating multidimensional intervention strategies aimed at enhancing resilience in this population. It is particularly essential to design integrated approaches that extend beyond individual-level interventions to encompass



community-based and policy-level strategies, thereby addressing the broader structural and social determinants of mental health and suicide risk. At the micro level, employing positive psychology and cognitive-behavioral strategies may enhance cognitive, emotional, and behavioral coping abilities in response to stressful situations, thereby strengthening resilience. This suggests that individual-level interventions focusing on psychological resources can play a pivotal role in suicide prevention among older adults living alone. At the macro level, more realistic policies are needed at the national level to ensure that older adults living alone can lead safe lives within their communities. Housing, health, and welfare policies are crucial, as are economic policies that can guarantee basic living standards. Building on these, we must establish a social structure that ensures adequate care within the community.

Second, this study focused on older adults aged 70 and above living alone who were receiving the government's specialized tailored care services—an identified high-risk group characterized by elevated levels of depression and suicidal ideation.

Therefore, psychological protective factors such as resilience identified in this study can serve as foundational data for evaluating the effectiveness of future specialized services and for developing related programs aimed at high-risk older adults living alone.

Meanwhile, this study has several limitations. First, the sample was limited to recipients of the national specialized care services in three selected local government areas, which restricts the generalizability of the findings.

Additionally, due to practical challenges in recruiting older adults living alone, data collection was conducted over a relatively short period. These factors necessitate caution in interpreting the results. Future research should consider expanding the survey area nationwide and allocating sufficient time to include samples with more diverse characteristics and backgrounds, thereby enhancing the reliability and validity of the study. Lastly, this study tested the mediating effect of resilience on depression and suicidal ideation among older adults based on a limited sample size ( $n=87$ ). However, it is necessary to further examine the moderated mediating effect of resilience on depression and suicidal ideation in later life, taking into account potential moderators not identified in this study, and to verify these effects through longitudinal research that captures changes over time.

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## AUTHOR CONTRIBUTIONS

Dr. Hee Jung KIM had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis. All authors reviewed this manuscript and agreed to individual contributions.

Conceptualization: all authors. Methodology: all authors. Investigation: all authors. Data curation: all authors. Formal analysis: all authors. Writing—original draft: HYK and HJK. Writing—review & editing: HJK.

## CONFLICTS OF INTEREST

No existing or potential conflict of interest relevant to this article was reported.

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None.

## DATA AVAILABILITY

The data presented in this study are available upon reasonable request from the corresponding author.

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# Associations between Fear of Cancer Recurrence and Health-Related Quality of Life, Psychological Distress, and Health Behaviors in Cancer Survivors

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## ABSTRACT

**Background:** Fear of cancer recurrence (FCR) is a prevalent psychological concern among cancer survivors, closely linked to diminished health-related quality of life (HRQoL) and increased psychological distress. However, its impact on health-related behaviors remains poorly understood. This study investigated associations between FCR, HRQoL, psychological distress, and health-related behaviors in a diverse cancer survivor population.

**Methods:** This cross-sectional study analyzed data from 326 adult cancer survivors across diverse cancer types. FCR was assessed using four items addressing concern about recurrence, metastasis, second cancer, and follow-up and participants were classified as high FCR. HRQoL was evaluated using the EuroQoL 5-Dimension Questionnaire (EQ-5D), EuroQoL Visual Analogue Scale (EQ-VAS), and European Organization for Research and Treatment of Cancer Quality of Life Questionnaire Core 30 (EORTC QLQ-C30); psychological symptoms were measured by the 9-item Patient Health Questionnaire (PHQ-9) and 7-item Generalized Anxiety Disorder scale (GAD-7). Health-related behaviors included recent screening, vaccination, and physical activity. Linear regression was used for continuous outcomes and logistic regression for binary outcomes, controlling for demographic and clinical covariates.

**Results:** Participants with high FCR demonstrated significantly poorer HRQoL across multiple domains, including Global Health Status/Quality of Life ( $\beta = -10.56$ , 95% confidence interval [95% CI]:  $-14.82$  to  $-6.29$ ,  $P < 0.001$ ), physical functioning ( $\beta = -7.04$ ,  $P < 0.001$ ), and emotional functioning ( $\beta = -12.86$ ,  $P < 0.001$ ). They also reported greater fatigue ( $\beta = 11.11$ ,  $P < 0.001$ ), pain ( $\beta = 9.85$ ,  $P < 0.001$ ), and financial difficulty ( $\beta = 17.32$ ,  $P < 0.001$ ). High FCR was strongly associated with depressive (PHQ-9:  $\beta = 3.10$ ,  $P < 0.001$ ) and anxiety symptoms (GAD-7:  $\beta = 2.70$ ,  $P < 0.001$ ). In contrast, no significant associations were found between FCR and health-related behaviors.

**Conclusions:** High FCR is strongly associated with impaired HRQoL and increased psychological distress but does not translate into increased engagement in preventive health behaviors. These findings underscore the need for integrated interventions addressing both emotional and behavioral aspects of survivorship care.

**Keywords:** Fear of cancer recurrence, Quality of life, Cancer survivors, Health behavior

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## INTRODUCTION

Advancements in cancer diagnosis and treatment, the expansion of national screening programs, and growing public awareness have led to a steady increase in the number of cancer survivors in Korea. As of 2020, approximately 2.56 million individuals in Korea had survived cancer, comprising nearly 5% of the national population, with the proportion of five-year survivors rising from 38.7% in 2010 to 57.7% in 2020 [1]. These trends signal a paradigm shift in the conceptualization of cancer—from an acute illness to a chronic condition requiring long-term survivorship care.

Even after completing primary treatment, many cancer survivors continue to face a variety of physical, psychological, and social challenges. Among these, fear of cancer recurrence (FCR) has emerged as one of the most common and persistent concerns [2-4]. FCR ranges from mild anxiety to clinically significant distress, with its intensity varying by cancer type, time since diagnosis, sex, and psychosocial characteristics. Excessive FCR has been linked to depression, anxiety, insomnia, and overall declines in health-related quality of life (HRQoL) [4,5].

FCR may also influence survivors' health behaviors, although the direction and magnitude of its effects remain unclear. For instance, a study of young breast cancer survivors revealed that individuals with high FCR were more likely to undergo screening but also engaged in detrimental behaviors such as smoking and alcohol use [6]. Other findings suggest that FCR can heighten somatic symptom sensitivity and psychological stress, influencing health behaviors and lifestyle management [7]. In colorectal cancer survivors, FCR has been associated with engagement in preventive behaviors like physical activity and smoking cessation [8]. Nevertheless, many prior studies have focused on narrow populations, limiting their generalizability. Moreover, some evidence indicates that while FCR significantly affects psychological functioning and HRQoL, its influence on health-related behaviors may be modest [9].

Accordingly, this study utilized data from the pilot phase of the Korean Nationwide Survey for Cancer Survivorship, the first population-based investigation of FCR in Korean survivors. By including six major cancer types (stomach, colorectal, liver, breast, cervical, and lung), the findings are more generalizable than those of cancer-specific studies. FCR was measured using a multidimensional tool covering recurrence, metastasis, secondary cancer, and follow-up concerns. This study assessed the associations between HRQoL, psychological symptoms, and

health behaviors, addressing the limitations of prior research focused on narrower populations or outcomes. These findings inform integrated survivorship care that considers both emotional and behavioral aspects.

## METHODS

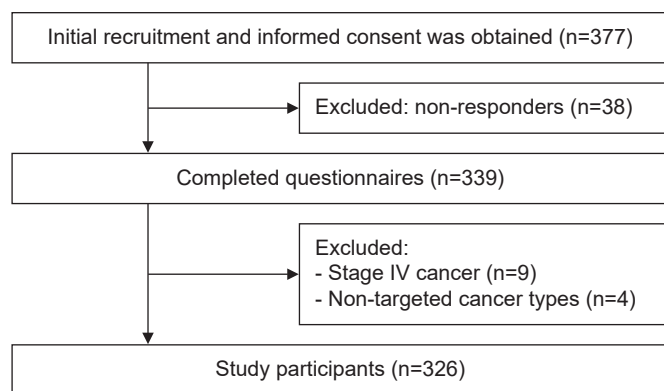
### Study design and participants

This study used data derived from a pilot survey conducted prior to the Korean Nationwide Survey for Cancer Survivorship. The pilot survey aimed to evaluate the feasibility of the study protocol, refine the questionnaire, and validate the data collection procedures. The survey covered a wide range of topics, including treatment history, current health status, quality of life, unmet needs, health behaviors, health policy indicators, and sociodemographic characteristics.

The eligibility criteria for the study participants were as follows: adults aged  $\geq 19$  years who were diagnosed with one of the six major cancers in Korea (stomach, colorectal, liver, breast, cervical, or lung cancer) and had survived for at least 1 year post-diagnosis. Participants were required to have completed primary cancer treatment, such as surgery, chemotherapy, or radiotherapy. Exceptions were made for individuals who were currently receiving oral chemotherapy without evidence of disease progression, who received adjuvant hormonal therapy following breast cancer treatment, or who exhibited no disease progression after receiving local treatment for liver cancer. Individuals currently undergoing treatment for recurrent or metastatic cancer and those diagnosed with terminal stage cancer were excluded.

Recruitment was conducted in collaboration with the medical staff at four institutions in Korea—namely, a national cancer center, two major private hospitals, and a regional cancer center. A total of 377 participants were enrolled from 6 November 2023 to 29 March 2024. Data were collected by a professional survey agency via either in-person interviews or self-administered online questionnaires, based on participant preferences. Out of 377 recruited individuals, 339 completed the survey. After excluding nine participants with stage IV cancer and four individuals diagnosed with non-major cancers (e.g., thyroid cancer, carcinoid sarcoma), 326 participants were finally included in the analysis (Fig. 1).

The study protocol was approved by the Institutional Review Board of the Chonnam National University Hwasun Hospital (IRB no. CNUHH-2025-073). Informed consent was obtained



**Fig. 1.** Flow diagram of participant.

through face-to-face interviews with trained research nurses using a web-based form.

### Covariates

Sociodemographic variables included age, sex, marital status, educational level, monthly household income, comorbidities, and employment status. Marital status was dichotomized into “single (including divorced or widowed)” and “married or cohabiting.” Educational level was categorized into three groups: middle school graduate or lower, high school graduate, and college graduate or higher. Monthly household income was divided into quartiles. Comorbidity status was assessed based on the presence of major chronic conditions and categorized into “none” and “one or more.” Employment status was classified according to current economic activity as either “employed” or “unemployed.”

Cancer-related clinical variables included cancer type, stage, treatment history, and time since diagnosis. Cancer types were classified as breast, stomach, colorectal, lung, cervical or uterine, and liver. Cancer stage was categorized as stage 0, I, II, or III or unknown. Treatment history included surgery, chemotherapy, radiotherapy, hormone therapy, and other procedures, each treated as a separate binary variable. Time since diagnosis was categorized into “1.0 to <4.5 years” and “4.5 to ≤10 years.”

FCR was assessed using four items that evaluated concerns regarding regular medical check-ups, the development of a second primary cancer, cancer recurrence, and metastasis. Each item was rated on a 4-point Likert scale ranging from 1 (“not at all”) to 4 (“very much”). To classify the level of FCR, the number of items rated as 3 (“quite”) or 4 (“very much”) was counted. Participants who endorsed a score of 3 or higher on two or more items were categorized as having high FCR, where-

as those with fewer than two such responses were classified as having low FCR. This classification reflects a multidimensional approach to capturing clinically meaningful levels of recurrence-related fear across multiple domains [10,11].

### Health-related quality of life and psychological distress

HRQoL was assessed using the EuroQol 5-Dimension Questionnaire (EQ-5D), EuroQol Visual Analogue Scale (EQ-VAS), and European Organization for Research and Treatment of Cancer Quality of Life Questionnaire Core 30 (EORTC QLQ-C30, version 3.0).

The EQ-5D [12] assesses five dimensions (mobility, self-care, usual activities, pain/discomfort, and anxiety/depression), each rated on a 3-point scale. For analysis, responses were dichotomized: “no problems” (score=1) coded as 0, and “some” or “extreme problems” (scores=2 or 3) coded as 1. The EQ-VAS [12] is a 20-cm vertical scale ranging from 0 (worst health) to 100 (best health), reflecting self-rated current health. The EORTC QLQ-C30 [13] is a 30-item questionnaire developed to measure HRQoL in patients with cancer. It comprises the three components: the five functional scales, the nine symptom scales, and a Global Health Status/Quality of Life (GHS/QoL) scale. Each item is rated on a 4-point Likert scale (“not at all” to “very much”) or a 7-point scale for GHS/QoL items. All scores were linearly transformed to a 0–100 scale, with higher scores indicating better functioning on the functional and GHS/QoL scales or worse symptoms on the symptom scales. The Korean version shows good reliability (Cronbach’s  $\alpha > 0.70$  for most subscales;  $\alpha = 0.60$  for cognitive function) [14].

Depression was measured using the Korean version of the 9-item Patient Health Questionnaire (PHQ-9) [15], with scores ranging from 0 to 27 and higher scores indicating more severe symptoms. The Korean version has high internal consistency ( $\alpha = 0.88$ ) [16]. Anxiety was assessed via the 7-item Generalized Anxiety Disorder scale (GAD-7) [17], with scores ranging from 0 to 21, also showing excellent reliability in Korean validation studies ( $\alpha = 0.91$ ) [18].

### Health-related behaviors

Health-related behaviors were assessed via self-administered questionnaires and categorized as change-or status-based.

Change-based behaviors included weight control, alcohol consumption, and smoking. Weight-control efforts were assessed using self-reported efforts to reduce, maintain, or gain weight over the past year. Alcohol consumption and smoking



behaviors were categorized as no change, reduction, cessation after diagnosis, or lifelong abstinence. Status-based behaviors included recent health check-ups (within 2 years), cancer screening (for cancers other than the index cancer), vaccinations (pneumococcal, herpes zoster, and seasonal influenza), physical activity, use of complementary and alternative medicine (CAM), and nutritional supplements.

Physical activity was assessed by the number of days per week participants walked  $\geq 10$  minutes and did strength training (e.g., push-ups, sit-ups, weightlifting) over the past week, categorized as  $< 5$  or  $\geq 5$  days for walking and  $< 2$  or  $\geq 2$  days for strength training [19]. CAM use was assessed as ever or never used for cancer treatment after the diagnosis. The current use of nutritional supplements, including vitamins and traditional remedies, was also recorded.

### Statistical analyses

Descriptive statistics are summarized as means with standard deviations for continuous variables and frequencies with percentages for categorical variables. Group comparisons based on FCR were performed using independent t-tests and chi-square tests.

To assess association between FCR and HRQoL or psychological symptoms, multivariate linear and logistic regression models were used for continuous and binary outcomes, respectively. All models were adjusted for age, sex, cancer type, cancer stage, comorbidities, time since cancer diagnosis, educational level, monthly household income, and employment status. Multicollinearity was assessed using variance inflation factors; all were  $< 1.45$ . Bonferroni correction was applied for multiple comparisons with a significance threshold of  $P < 0.0024$ .

Associations between FCR and health-related behaviors were examined using the chi-squared test. Multivariable logistic regression analyses assessed independent associations between each behavior and high FCR (coded 1), adjusting for demographic and clinical covariates. Each behavior was entered as the main independent variable in a separate model, and statistical significance was set at  $P < 0.05$ .

All analyses were conducted using Python version 3.11.4 (Python Software Foundation) [20], with the statsmodels (v0.14.0) [21], pandas (v2.0.3) [22], and scikit-learn (v1.3.0) [23] libraries.

## RESULTS

### Characteristics of the participants

A total of 326 cancer survivors were included in the analysis (Table 1), of whom 149 (45.7%) were classified as having high FCR and 177 (54.3%) as having low FCR. The mean age of participants was  $57.5 \pm 10.7$  years, with the high FCR group being younger than the low FCR group ( $56.3 \pm 10.2$  years vs.  $58.5 \pm 11.0$  years). Female accounted for 69.9% of the total sample and were slightly more represented in the high FCR group (71.1% vs. 68.9%). The majority of participants were married or cohabiting (73.6%) and had attained a college-level education or higher (55.2%). The distribution of monthly household income was relatively even across quartiles. Comorbid conditions were present in 30.1% of participants, and 56.1% were employed at the time of the survey. Breast cancer was the most common cancer type (32.2%), followed by stomach (23.0%) and colorectal cancer (15.0%). Most participants had been diagnosed with early-stage disease (stage 0–II, 71.8%). Nearly all participants had undergone surgery (94.5%), and over half had received chemotherapy (49.4%) or radiotherapy (39.6%). Time since diagnosis ranged from 1.0 to 10 years, with 53.4% having been diagnosed within the past 4.5 years. The mean total FCR score was  $9.9 \pm 3.3$ .

There were no statistically significant differences in sociodemographic or clinical characteristics between the low and high FCR groups (all  $P > 0.05$ ), except for cancer type ( $P = 0.005$ ). Specifically, breast cancer was more prevalent among participants with high FCR (43.6% vs. 22.6%), while colorectal cancer was more common in the low FCR group (27.1% vs. 18.1%). The distributions of other cancer types showed no clear differences between groups.

### Health-related quality of life and psychological symptoms according to fear of cancer recurrence

Table 2 presents the comparisons of HRQoL and psychological symptoms between participants with low and high FCR. After adjusting for age, sex, cancer type, cancer stage, comorbidity, time since diagnosis, education level, household income, employment status, and study group, several significant differences were observed across multiple domains.

On the EQ-5D, participants in the high FCR group were more likely to report problems in pain/discomfort (odds ratio [OR] = 2.31, 95% confidence interval [95% CI]: 1.44–3.71;  $P < 0.001$ ), and anxiety/depression (OR = 4.26, 95% CI: 2.58–7.02;  $P < 0.001$ ). The EQ-VAS score was also significantly lower in the



**Table 1.** Baseline characteristics of participants according to FCR status

Characteristic	Overall (n=326)	Low FCR (n=177)	High FCR (n=149)	P-value
Age (yr)	57.5±10.7	58.5±11.0	56.3±10.2	0.061
Sex				0.754
Male	98 (30.1)	55 (31.1)	43 (28.9)	
Female	228 (69.9)	122 (68.9)	106 (71.1)	
Marital status				0.225
Single	86 (26.4)	52 (29.4)	34 (22.8)	
Married and cohabiting	240 (73.6)	125 (70.6)	115 (77.2)	
Educational level				0.529
Middle school graduate or lower	34 (10.4)	19 (10.7)	15 (10.1)	
High school graduate	112 (34.4)	56 (31.6)	56 (37.6)	
College graduate or higher	180 (55.2)	102 (57.6)	78 (52.3)	
Monthly household income <sup>a</sup>				0.631
1Q	84 (25.8)	41 (23.2)	43 (28.9)	
2Q	80 (24.5)	44 (24.9)	36 (24.2)	
3Q	91 (27.9)	50 (28.2)	41 (27.5)	
4Q	71 (21.8)	42 (23.7)	29 (19.5)	
Comorbidity				0.754
0	228 (69.9)	122 (68.9)	106 (71.1)	
≥1	98 (30.1)	55 (31.1)	43 (28.9)	
Employment status				0.975
Unemployed	143 (43.9)	77 (43.5)	66 (44.3)	
Employed	183 (56.1)	100 (56.5)	83 (55.7)	
Cancer type				0.005
Breast	105 (32.2)	40 (22.6)	65 (43.6)	
Stomach	75 (23.0)	30 (16.9)	19 (12.8)	
Colorectal	49 (15.0)	48 (27.1)	27 (18.1)	
Lung	41 (12.6)	25 (14.1)	16 (10.7)	
Cervical or uterine	31 (9.5)	15 (8.5)	10 (6.7)	
Liver	25 (7.7)	19 (10.7)	12 (8.1)	
Cancer stage				0.140
0	24 (7.4)	17 (9.6)	7 (4.7)	
I	132 (40.5)	68 (38.4)	64 (43.0)	
II	78 (23.9)	38 (21.5)	40 (26.8)	
III	73 (22.4)	40 (22.6)	33 (22.1)	
Unknown	19 (5.8)	14 (7.9)	5 (3.4)	
Cancer treatment				
Surgery (yes)	308 (94.5)	168 (94.9)	140 (94.0)	0.894
Chemotherapy (yes)	161 (49.4)	83 (46.9)	78 (52.3)	0.384
Radiotherapy (yes)	129 (39.6)	58 (32.8)	71 (47.7)	0.009
Hormone therapy (yes)	60 (18.4)	22 (12.4)	38 (25.5)	0.004
Procedure (yes)	8 (2.5)	3 (1.7)	5 (3.4)	0.544
Time since diagnosis (yr)				0.651
1.0 to <4.5	174 (53.4)	97 (54.8)	77 (51.7)	
4.5 to ≤10	152 (46.6)	80 (45.2)	72 (48.3)	
FCR total score (range, 4–16)	9.9±3.3	7.4±1.4	12.9±2.1	<0.001

Values are presented as mean±standard deviation or number (%). FCR was assessed using four items addressing concerns about (1) routine check-ups, (2) the development of a second primary cancer, (3) recurrence of the original cancer, and (4) cancer metastasis. Each item was rated on a 4-point Likert scale (1="not at all" to 4="very much"). High FCR was defined as endorsement of a score ≥3 ("quite a bit" or "very much") on at least two of the four items, reflecting clinically meaningful levels of concern across multiple domains.

FCR, fear of cancer recurrence.

<sup>a</sup>Monthly household income was divided into quartiles.

**Table 2.** Comparison of health-related quality of life and psychological symptoms according to FCR status

	Range	Overall (n=326)	Low FCR (n=177)	High FCR (n=149)	Effect <sup>c</sup> (β or OR [95% CI])	P-value
EQ-5D						
Mobility	0–1	49 (15.0)	21 (11.9)	28 (18.8)	2.03 (1.03 to 3.99)	0.041
Selfcare	0–1	18 (5.5)	9 (5.1)	9 (6.0)	1.20 (0.45 to 3.25)	0.716
Usual activity	0–1	75 (23.0)	29 (16.4)	46 (30.9)	2.34 (1.33 to 4.11)	0.003
Pain/discomfort	0–1	166 (50.9)	72 (40.7)	94 (63.1)	2.31 (1.44 to 3.71)	<0.001 <sup>f</sup>
Anxiety/depression	0–1	164 (50.3)	63 (35.6)	101 (67.8)	4.26 (2.58 to 7.02)	<0.001 <sup>f</sup>
EQ-VAS <sup>a</sup>	0–100	70.3±17.6	74.3±16.2	65.6±18.0	-8.66 (-12.55 to -4.77)	<0.001 <sup>f</sup>
EORTC QLQ-C30 <sup>b</sup>						
GHS/QoL <sup>a</sup>	0–100	65.0±19.5	70.2±17.5	58.9±19.9	-10.56 (-14.82 to -6.29)	<0.001 <sup>f</sup>
Functional scales <sup>a</sup>						
Physical functioning	0–100	77.7±17.2	81.4±15.5	73.3±18.0	-7.04 (-10.80 to -3.27)	<0.001 <sup>f</sup>
Role functioning	0–100	77.6±24.1	82.1±23.7	72.1±23.6	-9.16 (-14.52 to -3.80)	<0.001 <sup>f</sup>
Emotional functioning	0–100	72.4±21.7	79.3±18.0	64.2±22.9	-12.86 (-17.39 to -8.33)	<0.001 <sup>f</sup>
Cognitive functioning	0–100	74.2±21.1	80.0±15.7	67.3±24.4	-9.82 (-14.24 to -5.40)	<0.001 <sup>f</sup>
Social functioning	0–100	74.3±25.5	81.3±21.5	66.0±27.4	-13.41 (-18.89 to -7.93)	<0.001 <sup>f</sup>
Symptom scales <sup>c</sup>						
Fatigue	0–100	38.3±23.0	32.5±20.8	45.3±23.6	11.11 (6.10 to 16.13)	<0.001 <sup>f</sup>
Nausea and vomiting	0–100	11.5±17.3	9.4±16.9	13.9±17.6	4.22 (0.27 to 8.16)	0.036
Pain	0–100	21.8±23.7	15.9±21.3	28.7±24.6	9.85 (4.78 to 14.91)	<0.001 <sup>f</sup>
Dyspnea	0–100	20.9±24.7	18.1±23.6	24.2±25.7	5.36 (-0.27 to 10.98)	0.062
Insomnia	0–100	33.7±31.5	25.4±27.8	43.6±32.9	14.51 (7.85 to 21.16)	<0.001 <sup>f</sup>
Appetite loss	0–100	16.0±22.6	13.4±21.4	19.0±23.7	6.01 (0.84 to 11.17)	0.023
Constipation	0–100	25.1±27.2	22.6±25.2	28.0±29.3	5.44 (-0.84 to 11.72)	0.089
Diarrhea	0–100	20.9±24.8	20.7±24.8	21.0±24.9	3.23 (-2.14 to 8.59)	0.238
Financial difficulties	0–100	25.9±29.4	17.1±22.8	36.2±32.9	17.32 (11.19 to 23.44)	<0.001 <sup>f</sup>
PHQ-9 <sup>d</sup>	0–27	5.4±5.3	3.8±3.9	7.3±6.0	3.10 (1.99 to 4.22)	<0.001 <sup>f</sup>
GAD-7 <sup>d</sup>	0–21	3.9±3.9	2.5±3.0	5.6±4.2	2.70 (1.89 to 3.52)	<0.001 <sup>f</sup>

Values are presented as number (%) or mean±standard deviation unless stated otherwise.

CI, confidence interval; EORTC QLQ-C30, European Organization for Research and Treatment of Cancer Quality of Life Questionnaire Core 30; EQ-VAS, EuroQol Visual Analogue Scale; EQ-5D, EuroQol 5-Dimension Questionnaire; FCR, fear of cancer recurrence; GAD-7, 7-item Generalized Anxiety Disorder scale; GHS/QoL, Global Health Status/Quality of Life; OR, odds ratio; PHQ-9, 9-item Patient Health Questionnaire.

<sup>a</sup>Higher scores indicate better health status or quality of life. <sup>b</sup>All scores were linearly transformed to a 0–100 scale according to the EORTC QLQ-C30 scoring manual. <sup>c</sup>Higher scores indicate greater symptom burden. <sup>d</sup>Higher scores indicate greater psychological symptom severity (depression or anxiety).

<sup>e</sup>Estimated using linear regression or logistic regression analysis after adjusting for age, sex, cancer type, cancer stage, comorbidity, time since cancer diagnosis, education level, monthly household income and job status. <sup>f</sup>Bonferroni correction was applied for multiple comparisons with a significance threshold of  $P<0.0024$ .

high FCR group ( $\beta=-8.66$ , 95% CI: -12.55 to -4.77;  $P<0.001$ ).

In the EORTC QLQ-C30 domains, the high FCR group reported significantly lower scores in all functional scales, including physical ( $\beta=-7.04$ ;  $P<0.001$ ), role ( $\beta=-9.16$ ;  $P<0.001$ ), emotional ( $\beta=-12.86$ ;  $P<0.001$ ), cognitive ( $\beta=-9.82$ ;  $P<0.001$ ), and social functioning ( $\beta=-13.41$ ;  $P<0.001$ ), as well as in GHS/QoL ( $\beta=-10.56$ ;  $P<0.001$ ). Regarding symptom burden, participants with high FCR had significantly higher levels of fatigue ( $\beta=11.11$ ;  $P<0.001$ ), pain ( $\beta=9.85$ ;  $P<0.001$ ), insomnia ( $\beta=14.51$ ;  $P<0.001$ ), and financial difficulties ( $\beta=17.32$ ;  $P<0.001$ ). Although nausea/vomiting and appetite loss were more prevalent in the high FCR group ( $P=0.036$  and  $P=0.023$ , respectively), they did not meet the Bonferroni-adjusted significance threshold ( $P<0.0024$ ).

Psychological symptoms were also markedly elevated among participants with high FCR. Both PHQ-9 ( $\beta=3.10$ , 95% CI: 1.99–4.22;  $P<0.001$ ) and GAD-7 scores ( $\beta=2.70$ , 95% CI: 1.89–3.52;  $P<0.001$ ) were significantly higher compared to the low FCR group.

### Health-related behaviors according to fear of cancer recurrence

Table 3 presents both the unadjusted prevalence of health-related behaviors in the low- and high-FCR groups and the adjusted associations derived from multivariable logistic regression. In chi-square analyses, no statistically significant differences were observed between groups for any change-based or status-based behavior. After adjusting for age, sex, cancer type, cancer stage,

**Table 3.** Comparison of health-related behaviors according to FCR status

	Low FCR (n=177)	High FCR (n=149)	P-value	OR (95% CI) <sup>a</sup>	P-value
Change-based behaviors					
Weight control effort	139 (78.5)	121 (81.2)	0.645	1.00 (0.65–1.53)	0.998
Change in alcohol consumption					
No change	43 (24.3)	27 (18.1)	0.400	Reference	0.265
Abstinence after cancer diagnosis	72 (40.7)	66 (44.3)	0.400	1.41 (0.75–2.63)	0.287
Lifetime non-drinker	62 (35.0)	56 (37.6)	0.400	1.45 (0.75–2.77)	0.266
Change in smoking status					
Continued or reduced smoking	12 (6.8)	9 (6.0)	0.702	Reference	0.427
Smoking cessation after cancer diagnosis	43 (24.3)	31 (20.8)	0.702	0.97 (0.36–2.62)	0.954
Lifetime non-smoker	122 (68.9)	109 (73.2)	0.702	1.18 (0.48–2.91)	0.719
Status-based behaviors					
General health check-up	153 (86.4)	130 (87.2)	0.960	0.97 (0.61–1.56)	0.902
Cancer screening	97 (54.8)	96 (64.4)	0.099	1.21 (0.76–1.91)	0.418
Pneumococcal vaccination	69 (39.0)	51 (34.2)	0.440	1.00 (0.64–1.56)	0.985
Herpes zoster vaccination	61 (34.5)	49 (32.9)	0.855	0.97 (0.62–1.52)	0.902
Seasonal influenza vaccination	70 (39.5)	65 (43.6)	0.528	1.02 (0.66–1.58)	0.933
Walking ≥5 day/wk	126 (71.2)	105 (70.5)	0.984	0.97 (0.64–1.47)	0.897
Strength training ≥2 day/wk	57 (32.2)	38 (25.5)	0.229	1.02 (0.67–1.56)	0.936
CAM use	35 (19.8)	21 (14.1)	0.521	0.97 (0.55–1.71)	0.907
Nutritional supplement use	57 (32.2)	105 (70.5)	0.856	0.99 (0.60–1.65)	0.975

Values are presented as number (%).

CAM, complementary and alternative medicine; CI, confidence interval; FCR, fear of cancer recurrence; OR, odds ratio.

<sup>a</sup>Estimated using logistic regression analysis after adjusting for age, sex, cancer type, cancer stage, comorbidity, time since cancer diagnosis, education level, monthly household income and job status.

comorbidity, time since diagnosis, educational level, household income, and employment status, none of the health-related behaviors remained significantly associated with high FCR. For example, the adjusted OR for general health check-up was 0.97 (95% CI: 0.61–1.56;  $P=0.902$ ), for cancer screening 1.21 (95% CI: 0.76–1.91;  $P=0.418$ ), and for CAM use 0.97 (95% CI: 0.55–1.71;  $P=0.907$ ).

## DISCUSSION

This study examined how FCR is associated with HRQoL, psychological symptoms, and health-related behaviors among cancer survivors. While high FCR was clearly associated with greater psychological distress and significantly poorer HRQoL across multiple domains, it did not correspond to increased engagement in health behaviors such as screening, physical activity, or vaccination. These findings suggest a divergence between emotional burden and behavioral response, underscoring the need for integrated survivorship strategies that address both psychosocial and behavioral dimensions.

Consistent with the present study, prior research has shown that high FCR are significantly associated with diminished HRQoL and heightened psychological distress. Previous study

of colorectal cancer survivors reported that higher FCR levels were correlated with lower functional scores on the EORTC QLQ-C30 and increased symptoms of anxiety and depression [24]. Similarly, research involving Korean stomach cancer survivor populations has indicated that FCR in stomach cancer survivors was associated with social, psychological, and HRQoL factors [5]. Consistent findings have also been observed among patients with advanced cancer, wherein high FCR was associated with substantial psychosocial burden [25]. The present study aligns with this body of evidence, revealing that individuals with high FCR exhibited significantly poorer scores across all functional domains of HRQoL—including physical, role, emotional, cognitive, and social functioning—as well as higher levels of symptom burden, particularly in relation to fatigue, insomnia, pain, and financial difficulties.

In contrast to symptoms such as fatigue, pain, insomnia, and financial difficulties, gastrointestinal (GI) symptoms—including nausea/vomiting, appetite loss, dyspnea, constipation, and diarrhea—demonstrated no statistically significant association with FCR ( $P>0.0023$ ). GI symptoms are often perceived as non-specific or treatment-related sequela rather than indicators of recurrence, particularly among non-GI cancers [5,24]. Moreover, FCR is fundamentally a cognitive-emotional construct centered

on the concern that cancer may return; it is more tightly tied to psychological distress and diminished functional domains than to isolated somatic complaints [3,4]. These findings are consistent with emerging evidence indicating that FCR correlates more strongly with fatigue, pain, and anxiety than with GI symptoms. Such a discrepancy emphasizes the need to consider both the interpretation and perceived salience of symptoms in understanding the psychological correlates of FCR.

Notably, multiple measures assessing psychological symptoms were included in this study, namely the EQ-5D anxiety/depression item, PHQ-9, and GAD-7. While each assesses facets of psychological well-being, they differ in both conceptual scope and psychometric design. The EQ-5D offers a singular, general indicator of emotional status within the broader context of HRQoL, whereas the PHQ-9 and GAD-7 are standardized multi-item instruments specifically developed to evaluate the severity of depressive and anxiety symptoms, respectively. The concurrent use of these measures permits a complementary assessment of emotional functioning, capturing both general distress and disorder-specific symptomatology.

Given that high FCR was not associated with increased engagement in screening, vaccination, weight control, alcohol use, smoking, physical activity, or supplement intake, several interpretative considerations warrant discussion. Contrary to the “teachable moment” hypothesis, which posits cancer experience and related fear motivate preventive action, findings suggest emotional distress alone may be insufficient to drive behavioral change [26,27]. This highlights the need for interventions combining psychosocial support with structured behavior-change strategies, rather than relying solely on psychological or educational approaches [28,29]. The conceptual distinction between psychological distress and behavioral adherence underscores the value of integrated multidisciplinary survivorship care that addresses both domains. CAM use and supplement intake were not evidence-based behaviors, but reflected broader coping responses to FCR. Although not medically validated, these behaviors may reflect emotionally driven adaptations offering insight into how cancer fear manifests behaviorally [30–32]. Moreover, recent evidence suggests FCR may indirectly affect behavior through unmet supportive care needs linked to lower screening participation, emphasizing the importance of addressing structural and psychosocial barriers alongside distress [9]. Finally, longitudinal research is needed to clarify how changes in FCR relate to health behaviors and whether individuals with persistently high FCR respond better to tailored interventions,

providing more nuanced survivorship strategies [33,34].

In this study, FCR was not associated with cancer stage or time since diagnosis. While this may partly reflect limited variability, stage IV cases were excluded, and over half the participants were long-term survivors, possibly due to analytical constraints, as the time since diagnosis was dichotomized at 4.5 years. However, previous studies have reported high FCR in early stage [35,36] and long-term survivors [37,38], suggesting FCR is not solely determined by clinical status or time. It reflects the complex interplay of clinical and psychosocial factors, including cancer type, treatment history, and perceived support. Longitudinal studies are needed to clarify these associations with long-term survival, suggesting clinical indicators alone may be insufficient.

This study has several limitations. First, its cross-sectional design limits causal inferences. Although high FCR was associated with poor HRQoL and psychological distress, reverse causality is plausible; emotional distress may heighten FCR rather than result from it. Similarly, the null association with preventive behaviors should be interpreted cautiously, as distress may hinder engagement. These limitations highlight the need for longitudinal studies to clarify the directionality. Second, most variables were assessed via self-reports, potentially introducing recall and reporting biases. Social desirability may have influenced responses on psychological symptoms and health behaviors, possibly affecting the associations. Third, as a pilot study, the sample was relatively small, which may limit statistical power and generalizability. Fourth, FCR was assessed with four items on follow-up, secondary cancer, recurrence, and metastasis. High FCR was defined as a score  $\geq 3$  on at least two items, based on prior studies [39–41]. This approach balances clinical relevance and brevity, capturing key domains of recurrence-related fear while reducing respondent burden. Although not externally validated, it was deemed appropriate for a pilot study with a moderate sample size. Further refinement and psychometric validation are planned from an upcoming nationwide survivor survey. Fifth, inconsistent timeframes in assessing health-related behaviors may have affected the interpretation. Some behaviors (e.g., smoking, alcohol use, and weight control) were measured as post-diagnosis changes, whereas others (e.g., vaccination, physical activity, CAM use, and supplement intake) reflected current status. This heterogeneity complicates whether FCR affects initiation or maintenance of behaviors. Null findings should be interpreted cautiously; future studies should adopt standardized timeframes or longitudinal designs for clar-

ification.

Despite these limitations, this study had several strengths. It provides an integrated perspective on the relationships between FCR, HRQoL, psychological distress, and health-related behaviors in a diverse cohort of cancer survivors. Findings showed high FCR was strongly associated with poorer HRQoL and greater psychological burden, yet did not lead to increased participation in health-promoting behaviors, such as screening, vaccination, or physical activity. This discrepancy challenges the assumption that emotional distress—particularly fear—naturally motivates positive behavioral changes. Clinically, the results underscore the need for multidisciplinary survivorship programs addressing psychological and behavioral dimensions rather than relying on FCR alone as a driver of health behavior. Tailored interventions integrating emotional support with structured behavioral-change strategies may be essential for promoting long-term health among cancer survivors.

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Dr. Yu Ri CHOE full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy of the data analysis. All authors reviewed this manuscript and agreed to individual contributions.

Conceptualization: all authors. Data curation: YK and YRC. Formal analysis: SHC and YRC. Investigation: SHC. Methodology: YK and YRC. Project administration: YK and YRC. Software: SHC and YRC. Supervision: YRC. Validation: YK and YRC. Visualization: SHC and YRC. Writing—original draft: SHC. Writing—review & editing: SHC and YRC.

## CONFLICTS OF INTEREST

No existing or potential conflict of interest relevant to this article was reported.

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## DATA AVAILABILITY

The data presented in this study are available upon reasonable request from the corresponding author.

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## Correction to "Association between Caffeinated Beverages Consumption and Sleep Quality of Urban Workers"

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The authors would like to correct the errors in the publication of the original article (Korean J Health Promot 2025;25(1):9-19; <https://doi.org/10.15384/kjhp.2024.00101>). The first author's affiliation was corrected to "Department of Epidemiology and Health Promotion, Graduate School of Public Health, Yonsei University" and "Department of Medical Sciences, Graduate School of Ajou University", which also resulted in a change to the overall affiliation numbering.

The corrected details are given below for your reading.

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The authors apologize for any confusion or inconvenience this may have caused and appreciate your understanding.

# Instructions for Authors

Revised in January, 2024 (6th)  
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**Question:** Is intermittent high-dose vitamin D supplementation effective in the prevention of falls and fractures?

**Findings:** In this meta-analysis of 15 randomized controlled trials, intermittent high-dose vitamin D supplementation showed no beneficial effect in the prevention of falls and fractures and even showed a harmful effect in the high-quality trials.

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The submitting author is responsible for completing this information at submission, and it is expected that all authors will have reviewed, discussed, and agreed to their individual contributions ahead of this time.

### *<Example>*

Dr. MYUNG had full access to all of the data in the study and takes responsibility for the integrity of the data and the accuracy

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of the data analysis. All authors reviewed this manuscript and agreed to individual contributions.

Conceptualization: SKM.

Data curation: SWO and YJC.

Formal analysis: YJC.

Methodology: SKM, SWO, and YJC.

Software: SKM and YJC.

Writing - original draft: YJC.

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